Assessing the impacts of enriched information on crash prediction performance

Peer-reviewed author version

MARTINS GOMES, Monique; PIRDAVANI, Ali; BRIJS, Tom & Souza Pitombo, Cira
(2019) Assessing the impacts of enriched information on crash prediction
performance. In: Accident analysis and prevention, 122, p. 162-171.

DOI: 10.1016/j.aap.2018.10.004

Handle: <http://hdl.handle.net/1942/27270>

Assessing the impacts of enriched information on crash prediction performance

Monique Martins Gomes¹, Ali Pirdavani², Tom Brijs³, Cira Pitombo⁴

¹UHasselt, Transportation Research Institute (IMOB), Agoralaan, 3590 Diepenbeek, Belgium, and São Carlos School of Engineering, Univ. of São Paulo, Av. Trabalhador São Carlense, 400, São Carlos, Brasil (corresponding author). E-mail: moniqmartins@hotmail.com

²UHasselt, Faculty of Engineering Technology, Agoralaan, 3590 Diepenbeek, Belgium. E-mail: ali.pirdavani@uhasselt.be

³UHasselt, Transportation Research Institute, Agoralaan, 3590 Diepenbeek, Belgium. E-mail: tom.brijs@uhasselt.be

⁴São Carlos School of Engineering, Univ. of São Paulo, Av. Trabalhador São Carlense, 400, São Carlos, Brasil. E-mail: cirapitombo@gmail.com

Abstract:

While high road safety performing countries base their effective strategies on reliable data, in developing countries the unavailability of essential information makes this task challenging. As a result, this drawback has led researchers and planners to face dilemmas of “doing nothing” or “doing ill”, therefore restricting models to data availability, often limited to socio-economic and demographic variables. Taking this into account, this study aims to demonstrate the potential improvements in spatial crash prediction model performance by enhancing the explanatory variables and modelling casualties as a function of a more comprehensive dataset, especially with an appropriate exposure variable. This includes experimental work, where models based on available information from São Paulo, Brazil, and Flanders, the Dutch speaking area of Belgium, are developed and compared with each other. Prediction models are developed within the framework of Geographically Weighted Regression with the Poisson distribution of errors. Moreover, casualties and fatalities as the response variables in the models developed for Flanders and São Paulo, respectively, are divided into two sets based on the transport mode, called active (i.e., pedestrians and cyclists) and motorized transport (i.e., motorized vehicle occupants). In order to assess the impacts of the enriched information on model performance, casualties are firstly associated with all available variables for São Paulo and the corresponding ones for Flanders. In the next step, prediction models are developed only for Flanders considering all the available information in the Flemish dataset. Findings showed that by adding the supplementary data, reductions of 20% and 25% for motorized transport, and 25% and 35% for active transport resulted in AICc and MSPE, respectively. Considering the practical aspects, results could help identify hotspots and relate most influential factors, suggesting sites and data, which should be prioritized in future local investigations. Besides minimizing costs with data collection, it could help policy makers to identify, implement and enforce appropriate countermeasures.

Keywords: Crash prediction models, Geographically Weighted Regression, Road safety, Enriched data

1. Introduction

Despite efforts to improve road safety, an estimated 1.25 million victims of road crashes worldwide still die every year. Developing countries, which have low and middle incomes, account for 90 percent of this number (WHO, 2015), which is likely to rise even more if proper safety countermeasures and investments are not made.

Among the countries attempting to prevent road fatalities, Brazil alone is responsible for up to fifty thousand deaths and five hundred thousand injured people every year. They are casualties of over one million crashes per year in the country (DATASUS, 2014; WHO, 2015). As in other developing countries, this problem has been attributed to an insufficient development of supportive road infrastructure, policy changes and enforcement which have not taken into account urban intensification and the steady increase in vehicle use. In spite of the growing awareness on the urgency to

reverse these trends and efforts toward road safety programs and campaigns, the country's performance remains below expectations leading to an exponential rise in the number of casualties.

Road safety has long been a priority in developed countries. In addition to ongoing efforts regarding the implementation and good practice of successful countermeasures involving infrastructure, vehicle and road user behavior, developed countries have invested a great amount of time and effort developing their road safety strategies at the planning level, for instance by collecting and making comprehensive sources of reliable data available. Specifically in Europe, these strategies at both safety-planning and operational levels have led to a steady reduction in the number of deaths, therefore allowing the European fatality rates to decrease far below the global average (9.3 per 100,000 population, relative to the global rate of 17.4) (WHO, 2015).

In this context, spatial Crash Prediction Models (CPM) are a critical component in terms of safety planning considering both prediction and impact analysis purposes. This argument is valid as CPMs enable the estimation of values while providing an insight of the spatially varying relationship between crashes and related factors. Hence, an appropriate set of potential explanatory variables is crucial. As suggested in the literature, it should include variables that have a major influence on the dependent variables in previous studies, which can be measured in a valid and reliable way, are not endogenous (Elvik, 2007), and above all, consider what people are exposed to that could result in a crash (Carroll, 1971; Chapman, 1973; Hauer, 1982; Hauer, 1995; Hauer et al, 1996; Stewart, 1998; Qin et al, 2004; de Guevara et al., 2004; Elvik, 2007), such as the fact that absence of an exposure variable can lead to biased results (Jovanis and Chang, 1986; Fristrøm et al., 1995; Miaou et al., 2003). In this respect, there are six major groups of influential factors, namely: human, vehicle related, road design, environmental, time and traffic related factors (Kononov, 2002; Valent et al., 2002; Yau, 2004; Shankar et al., 1995; Delen et al., 2006; Elvik, 2007).

Human factors are commonly associated to driver behavior, e.g., alcohol and drug use, negligent and careless vehicle operations, failure to properly use protection devices, using cell phones or texting while driving, fatigue, etc. (Petridou and Moustaki, 2000; Odgen, 1996; Redelmeier and Tibshirani, 1997; Movig et al., 2004). Vehicle related factors refer to the characteristics of the vehicle such as its safety design standards, i.e., active and passive vehicle safety systems (Harvey and Durbin, 1986; Robertson, 1996; Langley et al., 2000; Bédard et al., 2002; Richter et al., 2005). Traffic related volumes are commonly represented by the Average Annual Daily Traffic (AADT) or the Vehicle Miles Travelled (VMT), and are both commonly used as exposure variables in prediction models (Hauer, 1995; Zhou and Sisiopiku, 1997; Martin, 2002; Qin et al., 2004; Pei et al., 2012; Xu et al., 2012; Ahmed et al., 2012; Pirdavani et al., 2013a). Road design refers to the road geometry and roadside conditions, such as well-designed curves and grades, wide lanes, adequate sight distance, clearly visible striping, appropriate design speeds and road categorization, flared guardrails, roadsides free of obstacles,

well-located crash attenuation devices, and well-planned use of traffic signals (Miaou, 1994; Taylor et al., 2000; Amundsen and Ranes, 2000; Kloeden et al., 2001; Karlaftis and Golias, 2002; Nilsson, 2004; Aarts and van Schagen, 2006; Rengarasu et al., 2007). Environmental related factors are weather and light conditions, for example (Shankar et al., 1995; Andrey and Knapper, 2003; Golob and Recker, 2003; Ahmed et al., 2012; Brijs et. al., 2008). Finally, time factors are related to the season or month of the year, weekends or weekdays, or the time of crash occurrence (Doherty et.al, 1998; Qin et al., 2006; Hao et al., 2016).

In spite of the scientific and technological advances toward safety promotion, unfortunately there is a gap between research and practice. Especially in developing countries where data availability has been an issue, this has discouraged researchers and policy makers, as they often find themselves in situations where they have to choose between doing nothing or restricting CPM to data availability. However, even considering efforts made in associating crashes with the available explanatory variables in such circumstances, this drawback leads to modelling errors, and thus unreliable predictions. Besides not being statistically reliable, they fail in terms of impact analysis that could further help to implement appropriate safety countermeasures. One explanation could be the existence of omitted variable bias, which in particular plays an important role in CPM reliability, generating biased and inconsistent estimates and coefficient signs (Washington et al., 2010; Mitra and Washington, 2012).

In order to highlight the importance of a comprehensive set of explanatory variables within CPM, this study aims to assess the potential impacts of enriched information on model performance, scrutinizing their improvements in terms of statistical and practical contributions. This includes empirical work, where models based on crash-related available information from São Paulo, Brazil, and Flanders, the Dutch speaking area of Belgium, are developed and compared. CPMs are developed within the framework of Geographically Weighted Regression (GWR) with the Poisson distribution of errors (GWPR).

Subsequently, GWPR models followed by a sensitivity analysis allow us to identify the statistical contribution of all information that was entered in the prediction models. To the best of our knowledge, this practice has only been explored in terms of microscopic-level analysis, focusing on the influence of the Highway Safety Manual (HSM) data variables on safety predictions. Some approaches found in previous studies include the Fractional Factorial Method (Akgüngör and Yıldız, 2007), Boosted Regression Trees (BRT) (Saha et al., 2015) and the “change one-factor-at-a-time” approach, which is the most commonly used sensitivity method in literature (Alluri and Ogle, 2012; Findley et al., 2012; Jalayer and Zhou, 2013; Williamson et al., 2015). In the present study, this investigation is conducted by analytically adding each variable to the prediction models, altering the other ones and evaluating the statistical contribution of each variable one at a time and their interactions, thus accounting for simultaneous variation of the input variables.

2. Study area and data

Spatial CPMs were developed based on available information of crashes, road networks and socio-economic and demographic variables from the state of São Paulo and Flanders. In the models developed for São Paulo, this information was geographically aggregated to the centroids of 644 municipalities out of 645 that are comprised by the state. São Paulo city itself was not included in the analyses, given its atypical values, which are far higher than the ones for other cities. Flemish models were produced at a zonal level, comprising 2,198 Traffic Analysis Zones (TAZs). The average size of a TAZ is 6.09 square kilometers with a standard deviation of 4.78 kilometers, and an average number of inhabitants equal to 2,416 persons.

For both regions, casualties and fatalities as the response variables in the models developed for Flanders and São Paulo, respectively, were divided into two sets based on the transport mode, called active transport and motorized transport. Casualties/fatalities for active transport included pedestrians and cyclists, while for motorized transport they were associated with motorized vehicle occupants. Moreover, records from a period of three years were used to produce the dependent variable for São Paulo (2009-2011) and Flanders (2010-2012). In the Brazilian models, this information was gathered from the Mortality Information System (*Sistema de informações de Mortalidade – SIM*), which is a public source created by DATASUS (2014). For the Flemish models, this and the remaining information were provided by the Ministry of Mobility and Public Works (MOBIEL VLAANDEREN).

Other socio-economic and demographic information from São Paulo was gathered from the last census index of 2010, made available by the Brazilian Institute for Geography and Statistics (IBGE, 2014). Given the limitations to obtain road feature information in Brazil, we included a “link length” as a proxy variable of the road network. To this end, we used the available road network available in OPENSTREETMAP. For both regions, the link length for motorized transport included information concerning trunk, highway, primary, secondary and tertiary roads, as well as the link length of residential and living streets. For active transport, the same road features were implemented, although highways and respective trunk length information were replaced by cycle paths and link length information of other roads designed only for pedestrians (e.g., footway), according to the OPENSTREETMAP classification.

Tables 1 and 2 show a list of variables of all the variables collected for São Paulo and Flanders, respectively, together with their definition and descriptive statistics. Variables that were included in the final CPM developed for both regions of the study are marked in bold in the tables.

Table 1: Descriptive statistics of variables collected for São Paulo

	Variable	Description	Average	Min	Max	SD*
Fatality	Active transport	Total number of fatalities of active transport mode users over 3 years	7.34	0	295	23.428
	Motorized transport	Total number of fatalities of motorized transport mode users over 3 years	11.88	0	317	28.716
Network	Link length of active transport	Total link length of active transport in a municipality (km)	141.13	5.04	2626.15	210.95
	Link length of motorized transport	Total link length of motorized transport in a municipality (km)	153.04	5.29	2831.03	227.25
	Area	Total surface area in the municipality (km²)	383.03	5	1977	317.07
Socio-economic and demographic	Population	Total number of inhabitants in the municipality	46597.35	805	1221979	108465.83
	Male population	Total number of male inhabitants in the municipality	22902.55	422	595043	52538.13
	Female population	Total number of female inhabitants in the municipality	23694.81	383	626936	55938.55
	Population density	Total population per square kilometers per municipality	291.13	3.73	12519.10	1166.18
	AAGR	Average Annual Growth Rate 2000-2010 (%) in the municipality	1.03	-2.15	10.92	1.25
	Percentage male population	Percentage of male inhabitants in the municipality	50.52	45.76	81.09	2.52
	Percentage female population	Percentage of female inhabitants in the municipality	49.48	18.91	54.24	2.52
	Percentage proportion population	Rate between the number of men and woman in the municipality	102.97	84.36	428.86	17.88
	Urban population	Total number of inhabitants in the urban zone of the municipality	44150.48	627	1221979	107468.51
	Rural population	Total number of inhabitants in the rural zone of the municipality	2446.88	0	46284	3609.38
	HDI	Human Development Index	0.739	0.639	0.862	0.032
	GNP	Gross National Product	22501.11	7131.54	287646.17	18418.14
	Employed people	Total number of inhabitants with income	12678.37	155	405980	35725.41
	Occupied people	Total number of inhabitants who perform some activity (with income or not)	14931.77	211	471267	41144.02
Vehicle fleet	Motorcycle	Total fleet of motorcycles and tricycles	4744.68	24	100831	10938.16
	Microbus	Total fleet of microbuses	90.76	0	3544	264.87
	Car	Total fleet of cars	13536.09	133	487044	38052.31
	Truck	Total fleet of trucks	705.21	11	18144	1544.29
	Bus	Total fleet of buses	135.84	3	4445	330.34
	Total number of vehicles	Total number of vehicles	19212.58	220	612097	50296.09
Fuel consump.**	Gasoline	Total gasoline consumption	7961187.11	0	256246033	21723939.41
	Diesel oil	Total diesel oil consumption	15343179.63	0	295769873	32673917.02
	Fuel oil	Total fuel oil consumption	822438.64	0	44127640	3078410.70
	GLM	Total liquefied petroleum gas consumption	2304087.98	0	62823861	5948082.76
	Ethanol	Total ethanol consumption	9746540.07	0	342168947	25378940.38

*SD: Standard deviation; **Fuel consumption in liters

Table 2: Descriptive statistics of variables collected for Flanders

	Variable	Description	Average	Min	Max	SD*
Crash	Active transport	Total number of casualties of active transport mode users observed in a TAZ over 3 years	15.04	0	298	25.06
	Motorized transport	Total number of casualties of motorized transport mode users observed in a TAZ over 3 years	45.36	0	500	53.83
Network	Capacity	Hourly average capacity of links in a TAZ (Passenger car per direction/h)	1790.10	1200	7348	554.60
	Link length of active transport	Total link length of active transport in a TAZ (km)	14.85	0	88	10.31
	Link length of motorized transport	Total link length of motorized transport in a TAZ (km)	15.87	0.39	87.95	10.80
	Intersection density	Number of intersections per square kilometer	1.75	0	50.63	3.37
	Speed	Average speed limit in a TAZ (km/h)	69.40	31	120	10.91
	Area	Total surface area of a TAZ (km²)	6.09	0	45	4.78
	Link density	Total link length in a TAZ (km²)	3.37	0	20.44	2.41
	Intersection	Total number of intersection in a TAZ	5.80	0	40	5.90
	Highway	Presence of a highway in a TAZ, described as: “No” represented by 0 and “Yes” by 1		0	1	
	Urban	Is the TAZ in the urban area? “No” represented by 0 and “Yes” by 1		0	1	
	Suburban	Is the TAZ in the suburban area? “No” represented by 0 and “Yes” by 1		0	1	
Exposure	Number of trips of active transport	Average daily number of trips originating/destined from/to a TAZ involving active mode	1103.40	0	8630	1316.12
	Number of trips of motorized transport	Average daily number of trips originating/destined from/to a TAZ involving motorized transport	2750.09	0	22650	2642.17
	Vehicle Kilometers travelled (VKT) - Highway	Total VKT on highways in a TAZ	27471.82	0	946153	84669.53
	VKT - Other roads	Total VKT on roads other than highways in a TAZ	26662.85	0	303238	28133.04
Socio-economic and demographic	Car ownership	Car ownership per household in a TAZ	1.13	0	14.00	0.47
	School children	Total number of children living in a TAZ that attend some school	364.09	0	92.45	772.59
	Population	Total number of inhabitants in a TAZ	2614.53	0	15803	2582.60
	Households	Total number of households in a TAZ	1091.15	0	8062	1177.90
	Employees	Total number of employed people in a TAZ	888.73	0	16286	1575.31
	Income level	Average income of residents in a TAZ described as below:				
		“Monthly salary less than 2249 euro” represented by 0		0	1	
		“Monthly salary more than 2250 euro” represented by 1		0	1	

*SD: Standard deviation

3. Methodology

Geographically Weighted Regression

The spatially varying impacts of different risk factors across the study areas were explored within the framework of the local modelling approach, Geographically Weighted Regression (GWR) (see Brunsdon et al., 1996, Fotheringham et al., 1996, Fotheringham et al., 1997; Fotheringham et al., 2002) using the GWR 4.0 software package (Nakaya et al., 2005). Given that the number of casualties and fatalities as the response variables were the count data with discrete and non-negative integer values, GWR models were performed using the Poisson distribution error (GWPR).

Developed by Fotheringham and Brunsdon (Fotheringham et al., 2002), GWR models intend to address the non-stationary relationship between variables found in Generalized Linear Models (GLM). These models capture this spatial variation by fitting a regression model at each sample point. The result of this process is a set of local spatial parameters, described by Equation (1).

$$\ln[E(C)(l_i)] = \ln(\beta_0(l_i)) + \beta_1(l_i)\ln(Exposure) + \beta_2(l_i)x_1 + \dots + \beta_n(l_i)x_n \quad (1)$$

Where $E(C)$ is the expected crash frequency, β_0 , β_1 , β_2 and β_n are model parameters for determined location l_i , $Exposure$ is the exposure variable, and x_1 and x_n correspond to other explanatory variables.

One of the assumptions behind GWR is motivated by the First Law of Geography from Tobler (1970), which argues that “everything is related to everything else but closer things are more related to each other”. The closer the observed data is from the location of the parameter to be estimated, the greater the influence on the estimation of β at location i compared to those that are far from it. Hence, this influence is determined based on geographic weights, which are assigned in function of all neighboring observations using a kernel function (Fotheringham et al., 2002), e.g., Gaussian (Equation 2) and bi-square (Equation 3), which are the two most common choices of weighting schemes (Hadayeghi et al., 2010).

Gaussian function:

$$W_{ij} = e^{-0.5(\frac{w_{ij}}{b})^2} \quad (2)$$

Bi-square function:

$$W_{ij} = \begin{cases} (1 - (\frac{d_{ij}}{b})^2)^2 & \text{if } d_{ij} < b \\ 0 & \text{otherwise} \end{cases} \quad (3)$$

Where W_{ij} is the measure of contribution of location j when calibrating the model for location i . d_{ij} is the Euclidian distance between locations i and j and b is the bandwidth size defined by a distance metric measure.

In GWR, the bandwidth controls the size of the kernel (number of observations around each data point) and the rate at which weights decay with increasing distances. Thus, similar to the weighting scheme, the choice of the bandwidth size plays an important role in the performance of the GWR models, as it involves a trade-off between bias and variance. The size of the bandwidth is optimized either by distance (fixed kernel), or by the number of neighboring observations (adaptive kernel) (Fotheringham et al., 2002; Guo et al., 2008; Hadayeghi et al., 2010). In this study, GWR was performed using the bi-square form with an adaptive bandwidth, such that the bandwidth varies according to the data density, and the number of areas included in the kernel is kept constant.

Within different approaches to select the optimal bandwidth, minimizing cross validation (CV) or the corrected Akaike Information Criterion (AICc) are the most widely used (Pirdavani et al., 2014). However, while the former is given by the difference between observed and estimated values, the latter, additionally to the statistical Goodness-of-fit, rewards the complexity of the model, by imposing a penalty for increasing the number of estimated parameters (Fotheringham et al., 2002), expressed by the formulation in Equation (4).

$$AICc = D(b) + 2K(b) + 2 \frac{K(b)(K(b)+1)}{n-K(b)-1} \quad (4)$$

Where D and k denote the deviance and the effective number of parameters in the model with bandwidth b , respectively. Moreover, n denotes the number of observations. In this study, the bandwidth is calculated by means of AICc.

Methodological procedure

In order to demonstrate the potential improvements in the performance of spatial prediction models by enhancing the potential explanatory variables, the methodological procedure was divided into two main stages.

196 In the first stage, GWPR models were developed for São Paulo and Flanders, by only taking into account the same
197 explanatory variables available in both datasets. Given the limitations of the Brazilian dataset, the results of this stage
198 would reveal the best we could do with the available information of São Paulo, while there would be plenty of room yet
199 to improve the Flemish models. In the second stage, and in order to highlight the importance of having data which is as
200 complete as possible, GWPR models were developed for Flanders only, by considering all available variables in the
201 Flemish dataset.

202 At both stages, a multicollinearity test was conducted prior to the modelling steps, enabling us to select the most
203 significant variables to compose the final models. As a common practice, the Variance Inflation Factor (VIF) was used
204 to quantify how much the variance of the estimated regression coefficients increased if predictors were correlated. As a
205 common rule of thumb, 10 was defined as a cut off value, meaning that if VIF was higher than 10, then multicollinearity
206 was high (Kutner et al., 2004) and, therefore, these variables should not be present in the model simultaneously.
207 Subsequently, at the end of the first stage, models, for which we used the minimum data, were developed which were
208 called basic models.

209 In the second stage, the affluence of available information in the Flemish dataset enabled us to develop various distinct
210 models, and choose the one with the best overall fit. This exercise was conducted by having the intercept term as our
211 starting point and analytically combining variables with a VIF lower than 10. Hereafter, due to the greater complexity of
212 the GWR estimation procedure that conceivably causes interrelationships among local coefficient estimates when there
213 is no collinearity among the explanatory variables (Hadayeghi et al., 2010; Pirdavani et al., 2013b), at this stage, evidence
214 of multicollinearity among the produced local coefficient estimates was also observed. Hence, among all developed
215 CPMs, the best-fitted one was selected as it met the criteria of non-multicollinearity among variables and produced local
216 coefficient estimates, and subsequently based on the lowest AICc value. As a common rule-of-thumb, the difference
217 between the models was considered significant when the difference of AICc values between two models was higher than
218 4 (Charlton and Fotheringham, 2009). At the end of this stage, two models were developed for Flanders, which were
219 called improved models.

220 Finally, the performance of the improved models was compared with the basic models by means of AICc, Mean
221 Squared Prediction Error (MSPE), and the Pearson Correlation Coefficient (PCC).

222

223 *Sensitivity Analysis*

224 Flemish GWPR models followed by a sensitivity analysis helped us identify the statistical contribution of each
225 variable in the casualty prediction. In previous studies, this practice was commonly associated with micro-level analysis,

226 focusing on the influence of the HSM data variables on safety predictions. This exercise was held by altering the value
227 of one predictor variable at its maximum, minimum, and/or average, hence estimating the change in output relative to the
228 output generated from using the actual values of the variable. Therefore, the most influential variables were identified as
229 those that produced meaningful changes in the predicted values for the frequency of crashes (Saha et al., 2015).

230 In our study, this investigation was conducted by analytically adding each variable to the prediction models while
231 altering the other ones, and evaluating their statistical contribution by themselves and by their interactions, thus
232 accounting for simultaneous variations of the input variables. To this end, the intercept term was used as a starting point.
233 Hence, explanatory variables were analytically added to the prediction models and ranked according to their contribution
234 in terms of the model's performance. This contribution was measured by means of the AICc variations (%), where the
235 larger the reduction in AICc by the inclusion of a variable, the greater its contribution on the model performance.
236 Subsequently, this process was repeated with the remaining variables, but taking into account their interactions.
237 Thereupon, variables were tabulated according to their relative percentage of influence on the models in relation to the
238 intercept term, namely Relative I, and in relation to its previously best fitted model composition, namely Relative II.

239

240 **4. Results**

241 In the next sections, the results of the modelling practices for both case studies will be explained and discussed.
242 However, due to the limit of space in this paper, maps of significance and coefficient estimates are limited to motorized
243 transport in Flanders. Moreover, although it is not the main aim of this paper, we will briefly discuss the effect sizes in
244 relation to prior literature.

245

246 ***Results of modelling – Stage 1, Basic Models***

247 In spite of a great amount of available information in Brazil, most of it was limited to socio-economic and
248 demographic variables. As a result of this, most of the pieces of information collected were found to be correlated with
249 each other, therefore presenting high VIF values. Hence, produced basic models were limited to information concerning
250 the link length and population only. At this stage, the population was used as the exposure variable in its Natural
251 Logarithm (ln) form.

252 Table 3 shows the local parameter estimates of the basic GWPR models, for both dependent variables in São Paulo
253 and Flanders. This information is described by five number summaries: minimum, maximum (lower quartile, median
254 quartile and upper quartile), tabulated in this format and sequence.

Table 3: Local parameter estimates - Basic Models

Parameters	Active Transport		Motorized Transport	
	Brazil	Flanders	Brazil	Flanders
Intercept	-24.094, 15.716 (-13.546, -10.626, -8.081)	-8.890, 11.749 (0.153, 1.813, 2.966)	-16.228, 0.031 (-8.576, -6.837, -4.896)	-5.981, 8.582 (1.894, 3.336, 4.246)
Ln Population	-2.249, 2.463 (0.874, 1.155, 1.438)	-1.646, 1.668 (-0.046, 0.130, 0.365)	0.019, 1.898 (0.672, 0.871, 1.053)	-1.252, 1.307 (-0.063, 0.077, 0.267)
Link length	-0.006, 0.038 (-7.8e-04, 1.5e-04, 1.2e-03)	-0.264, 0.130 (-0.038, -0.012, 0.010)	-0.009, 0.010 (2.2e-04, 4.4e-04, 1.4e-03)	-0.149, 0.139 (-0.031, -0.008, 0.267)

Figure 1 shows maps of the local coefficient estimates as well as their significance at 0.05 level, for motorized transport in Flanders. In order to determine where relationships were significant (in blue) and where they were not (in brown), we computed the t-statistics. Thus, t-statistics between -1.96 and 1.96 were considered insignificant at 0.05 level, and values beyond this scale were considered significant at 0.05 level. Subsequently, Figure 2 shows maps with observed and predicted number of casualties.

Figure 1: Local coefficient estimates and significance maps (Motorized transport - Basic model - Flanders).

Figure 2: Observed and predicted number of casualties (Motorized transport - Basic model - Flanders).

Results of model performance at this stage revealed a negative correlation between casualties and both explanatory variables in a large number of TAZs, which in turn were significant at 0.05 confidence level. Given their rather direct association with exposure, we would expect that both explanatory variables were likely to have a positive correlation with casualties in most TAZs of this study, as also found in Wang et al. (2009 and Pirdavani et al. (2013b), for example.

271 One explanation for these counterintuitive signs could be that some variables, at some locations, are locally correlated
272 while no global multicollinearity is observed among them (Guo et al., 2008; Hadayeghi et al., 2010; Pirdavani et al.,
273 2014). Such local correlation has been attributed as the major reason for problems with counterintuitive signs. In order
274 to address this issue, at the second stage, we excluded variables at which high VIF values were found among the produced
275 coefficient estimates. The limitation of variables within the basic models restrained us from conducting this exercise at
276 the first stage. This drawback highlights the importance of a more diverse set of explanatory variables. Evidence in favor
277 of this assumption is shown in the second stage.

278 Such counterintuitive signs could also be a response to the omission of important variables, which leads to omitted
279 variable bias. Although not thoroughly explored in this study, one could assume the correlation of link length and other
280 road features or exposure variables that were omitted, therefore producing bias. This argument is valid as the exclusion
281 of essential variables, especially an exposure variable, could systematically invalidate further conclusions that could be
282 derived from the results (Washington et al., 2010; Mitra and Washington, 2012). A more in-depth investigation
283 concerning this problem could help to provide a better insight of the direction of these effects, which in this study, remains
284 speculative.

285

286 ***Results of modelling – Stage 2, Improved Models***

287 The Flemish dataset, in addition to significant information related to socio-economic, socio-demographic and road
288 networks (i.e., income level, speed, capacity, number of links, links and intersection density, presence of highways,
289 urbanization degree, to name a few) provided foremost diverse and suitable exposure variables, i.e., number of trips
290 (NOTs), vehicles flow and VKT. This enabled us to produce different models with different combinations of variables,
291 and choose the best fitted one.

292 After carrying out the VIF tests among variables and produced coefficient estimates, the final improved models for
293 active and motorized transport modes comprised the following information: NOTs, children attending school (school
294 children), road capacity (capacity), intersection density, car ownership, and VKT which was used as our exposure
295 variable. The respective coefficient estimates found for each explanatory variable are presented in Table 4 in the five
296 number summary format, and are followed in Figure 3 by their local coefficient estimates and significance maps.
297 Subsequently, Figure 4 presents the obtained maps for the observed and predicted number of casualties.

298

299 **Table 4: Local parameter estimates - Improved Models (Flanders)**

Parameters	Active Transport	Motorized Transport
Intercept	-14.835, 13.933 (-0.928, 1.580, 4.068)	-7.604, 11.041 (1.585, 3.312, 4.973)
NOTs	-0.002, 0.005 (-2.78-04, 4.8e-05, 4.4-04)	-9.4e-04, 0.001 (-1.04e-04, 2e-05, 1.23e-04)
School children	-0.007, 0.003 (-9.16-04, -2.33e-04, 3.4e-04)	-0.004, 0.003 (-5.74e-04, -1.43e-04, 2.33e-04)
Capacity	-0.008, 0.007 (-3.46e-04, 7.3e-05, 5.47e-04)	-0.007, 0.005 (-3.89-04, 1.5e-05, 4.5-04)
Intersection density	-4.183, 1.686 (-0.139, 0.023, 0.152)	-1.441, 1.171 (-0.086, 0.025, 0.131)
Ln VKT	-0.653, 1.486 (-0.054, 0.128, 0.306)	-0.666, 0.888 (-0.092, 0.076, 0.246)
Car Ownership	-7.378, 6.584 (1.649, -0.375, 0.445)	-6.763, 4.722 (-1.084, -0.128, 0.456)

302 **Figure 3: Local coefficient estimates and significance maps (Motorized transport - Improved model - Flanders).**

Figure 4: Observed and predicted number of casualties (Motorized Transport - Improved model - Flanders).

In general, results at this stage revealed a better performance in terms of overall model fit (see Table 5). From a statistical point of view, the improved models outperformed the basic models for both dependent variables. In the model developed for motorized transport, reductions of approximately 20% and 25% were observed compared to the basic model, for AICc and MSPE respectively. Likewise for the active mode, 25% and 35% reductions were obtained for AICc and MSPE, respectively.

Table 5: Comparison of parameters between basic and improved models (Flanders)

Parameters	Active Transport		Motorized Transport	
	Basic	Improved	Basic	Improved
GWPR AICc	29345.571	22553.607	60993.481	49525.611
Global AICc	50570.229	45486.734	102673.889	94634.071
MSPE	384.87	261.17	1786.66	1354.12
PCC	0.629	0.771	0.626	0.738

Moreover, results of a more diverse dataset, especially including an exposure variable, enabled the development of coherent coefficient estimates and signs. In this study, positive associations with casualties were found in most subzones, for intersection density, capacity, NOTs and VKT (likewise in Hadayeghi et al., 2003; Aguero -Valverde and Jovanis, 2006; Pirdavani et al., 2013a; Pirdavani et al., 2013b; Shariat-Mohaymany et al., 2015; Xu et al., 2017). On the contrary, casualties were found to have a negative correlation with school children and car ownership, in a large number of TAZs. This negative association with other social standing variables, i.e., income level, has been found in other previous studies (see Li et al., 2013; Pirdavani et al., 2013a; Pirdavani et al., 2013b; Pirdavani et al., 2014; Pirdavani et al., 2016), meaning that less casualties are expected to occur than in more affluent areas (in this study, where car ownership is higher). Given the negative association of casualties and school children, we would assume that speed limit and human factors associated to the driver's behavior, might have an influence.

In order to corroborate these assumptions and enable us to better understand the interactions between variables, such a macro-level analysis could be used as a basis for local investigations. This could help to enforce appropriate countermeasures, especially in areas where higher estimates were found. For instance, at subzones where casualties were found to have a positive association with school children, micro-level analysis could suggest changes in the speed limits

or signalizing intersections. This could be identified as the major contribution of having a more complete and diverse dataset. In spite of more reliable models, they would allow policy makers to prioritize subzones, and depending on the targets, specific TAZs could be used to investigate the interaction between variables, both within and outside of the models.

Subsequently, GWPR improved models followed by a step-wise approach helped us identify the statistical contribution of each variable in the crash prediction performance. The improvements in model performance by means of the percentage reductions found on AICc, for motorized transport, are shown in Figure 5. We considered the percentage difference by means of AICc in relation to the intercept term, namely Relative I, and in relation to its previously best fitted model composition, namely Relative II. In this illustration, we included part of the sensitivity analysis, for which capacity, for instance, showed the highest reduction by means of AICc in relation to the other variables, becoming the chosen variable with the intercept term. Subsequently, the improvements in model performance for active transport are shown in Table 6.

Model No.	Variables	AICc	Relative I (%)	Relative II (%)
1	Intercept	69996.71	-	-
2	Model 1 + capacity	65149.59	-6.92	-6.92
3	Model 2 + VKT	61120.22	-12.68	-6.18
4	Model 3 + school children	57616.19	-17.69	-5.73
5	Model 4 + intersection density	53909.95	-22.98	-6.43
6	Model 5 + car ownership	51521.46	-26.39	-4.43
7	Model 6 + NOTs	49525.61	-29.25	-3.87

Variables	AICc	%
Intercept	69996.71	-
Intercept + capacity	65149.59	-6.92
Intercept + VKT	65345.23	-6.65
Intercept + intersection density	57616.19	-5.61
Intercept + car ownership	53909.95	-5.57
Intercept + school children	51521.46	-5.56
Intercept + NOTs	67022.43	-4.25

Figure 5: Ranking of variables according to their contribution in the model for motorized transport.

Table 6: Ranking of variables according to their contribution in the model for active transport

Model No.	Variables	AICc	Relative I (%)	Relative II (%)
1	Intercept	33593.51	-	-
2	Model 1 + capacity	31360.08	-6.65	-6.65
3	Model 2 + car ownership	29132.76	-13.28	-7.10
4	Model 3 + intersection density	27310.70	-18.70	-6.25
5	Model 4 + school children	25276.7	-24.76	-7.45
6	Model 5 + NOTs	23876.5	-28.93	-5.54
7	Model 6 + VKT	22553.61	-32.86	-5.54

The results show that road capacity has the highest statistical contribution in the performance of CPM for active and motorized transport modes, suggesting that this information has priority over others. Secondly, VKT statistically contributes more to motorized transport models, while car ownership contributes more to the active transport models, and so on. This practice could be useful as it would help policy makers prioritize data collection, for instance by targeting variables that add higher statistical contributions to one specific travel mode or both, thus reducing costs with data collection.

5. Discussion and conclusions

The difficulty in obtaining crash-related information in Brazil, and its consequences in terms of model performance and development of potential studies that could help understand the crash phenomena and enforcement of appropriate countermeasures were the major reasons for carrying out this study. Although some data can be found in different road departments, police, health and census online sources (i.e., DATASUS, IBGE, DENATRAN), there is no link between their databases, and none of them are able to provide a full and effective data source with regard to accidents and fatalities in the country. Therefore, the absence of a comprehensive and complete database hampered the evaluation and follow-up of national road safety programs, as well as the development of studies that could contribute to national goals toward road safety.

Particularly concerning the explanatory variables collected in this study, most available information was restricted to socio-economic variables. In spite of their merits, socio-economic variables are often highly correlated with each other, and are not appropriate for safety-planning purposes, e.g., implementing safety countermeasures. Other developing countries have faced the same challenges, and this drawback has unfortunately led to the use of poor and unreliable CPMs to promote road safety in these countries.

366 In view of the above, this study aimed to demonstrate the potential improvements in the performance of spatial crash
367 prediction models by means of a more diverse dataset, including some potential explanatory variables. To this end,
368 benchmarking was carried out based on macro level CPMs developed with available fatality/crash-related information
369 from São Paulo and Flanders. In contrast to developing countries, European countries such as Belgium and other high
370 road safety performing countries have invested a great deal of time and money in obtaining crash-related information and
371 make them available through public channels and to academia. This practice has led to outcomes such as new strategies
372 and studies, and this trade-off has brought improvements in traffic safety by reducing the number of crashes, injuries and
373 fatalities.

374 Whereas the Brazilian dataset in this study was mostly limited to socio-economic variables, the Flemish dataset, in
375 addition to significant information related to socio-economic, socio-demographic and road networks, provided foremost
376 diverse and suitable exposure variables. As a result of this, improved models revealed lower values of AICc and MSPE,
377 for both dependent variables. Moreover, Flemish models at the second stage, presented a significant set of coefficient
378 estimates together with suitable coefficient signs. One potential outcome of the resulting macro-level CPMs could be the
379 identification of hotspots together with their major influence factors. Such results could be used as a reference to
380 microscopic investigations, and the implementation of suitable safety countermeasures' enforcement in a long term
381 transportation planning process.

382 Above all, this study addressed the strong dependence of CPMs on suitable and diverse input information, enabling
383 these models to perform as a "powerful tool", as is usually found in literature. Crashes are caused by multiple factors that
384 vary locally, and this complexity implies that ideally casualties are best predicted through a set of appropriate predictive
385 variables, including at least one potential exposure variable. By modelling casualties in Flanders based on the equivalent
386 available data in São Paulo, results were found not to be suitable. Apart from producing unreliable coefficient estimates,
387 they would also not be useful for safety planning and practical aspects. In other words, despite the efforts to improve the
388 statistical fit of the crash prediction models and associate crashes with the available explanatory variables in such
389 circumstances, these models fail to comprehensively explain the road casualties and, therefore, diminish the ability of
390 applying suitable safety countermeasures.

391 On the contrary, by modelling casualties based on the entire available data in Flanders, results revealed a better model
392 overall fit for both active and motorized transport modes. This suggests that a more diverse set of appropriate explanatory
393 variables, including a relevant exposure variable, is advantageous as it could address problems with counterintuitive signs
394 and omitted variable bias. In the ideal scenario, it could help policy makers determine local appropriate countermeasures

395 toward safety promotion (e.g., by altering the speed limits and intensifying local speed enforcement, improving
396 intersections, installing traffic management and control systems, implementing crosswalks, etc).

397 It is also worth mentioning that variables included in the Flemish models, besides having been found to be significant
398 in previous studies as well as in the present one, are less expensive and more accessible than those used in microscopic
399 analysis (e.g., driving data, braking and steering information or variables related to weather conditions). Moreover, they
400 are just examples of other potential information that could be used to develop those predictive/descriptive models.
401 Variables that are used in the models developed for Flanders could at most be interesting suggestions for extra data
402 collection in Brazil, as other local variables could also play a significant role, other than those included in the Flemish
403 models.

404 Last but not least, a sensitivity analysis was carried out allowing us to assess the statistical contribution of each
405 variable in the prediction model performance. Especially for countries where crash data is limited, either because of the
406 lack of financial resources or other imposed conditions, this practice could empower policy makers and responsible
407 offices to prioritize data collection. For instance, results revealed that data regarding road capacity would signify a major
408 statistical contribution for models, for both dependent variables. This is different from NOTs, for instance, which often
409 have priority in data collection, but as revealed in this study, would not bring such a significant contribution to the Flemish
410 models, neither for active, nor for motorized transport.

411 This investigation could have more value if similar analyses were carried out in different regions, based on their
412 available information. The consolidation of the produced results would enable, for instance, the development of a solid
413 benchmark, therefore, validating the priorities outlined in this study and helping to determine the importance of different
414 variables to model performance, in different areas.

415 In addition to this, for further studies, we suggest a more in-depth investigation addressing problems, such as
416 endogeneity and omitted variables to model performance, as they could help to verify the validity of the assumptions of
417 this study. One possible investigation could be for instance, to perform micro-level analysis in the identified hot zones,
418 therefore assessing model performance by adding and removing variables to the models.

419

420 **ACKNOWLEDGMENTS**

421

422 This research was supported by the Brazilian National Council for Scientific and Technological Development - CNPq.

423

424 **References**

425

426 Aarts, L. van Schagen, I., 2006. Driving speed and the risk of road crashes: a review. *Accid. Anal. Prev.* 38: p. 215-224.

427 Agüero-Valverde, J., Jovanis, P., 2006. Spatial Analysis of Fatal Injury Crashes in Pennsylvania. *Accid. Anal. Prev.* 38:

428 p. 618-25.

429 Ahmed, M., Abdel-Aty, M., Yu, R., 2012. Assessment of the interaction between crash occurrence, mountainous freeway

430 geometry, real-time weather and AVI traffic data. *Transportation Research Record.* 2280: p. 51-59.

431 Akgüngör, A. Yıldız, O, 2007. Sensitivity Analysis of an Accident Prediction Model by the 15 Fractional Factorial

432 Method. . *Accid. Anal. Prev.* 39 (1), p. 63-68.

433 Alluri, P., Ogle, J., 2012. Effects of state-specific SPFs, AADT estimations, and overdispersion parameters on crash

434 predictions using SafetyAnalyst. *Proceedings of the 91st Annual Meeting of the Transportation Research Board,*

435 *Washington, DC.*

436 Amundsen, F. H., Ranæs, G., 2000. Studies on traffic accidents in Norwegian road tunnels. *Tunnelling and Underground*

437 *Space Technology,* 15, 3-11.

438 Andrey, J., Knapper, C.K., 2003. *Weather and Transportation in Canada.* Department of Geography Publication Series,

439 no. 55. ISBN: 0-921083-65-3.

440 Bédard, M., Guyatt, G.H., Stones, M.J., Hirdes, J.P., 2002. The independent contribution of driver, crash, and vehicle

441 characteristics to driver fatalities. *Accid. Anal. Prev.* 34: p. 717-727.

442 Brijs, T., Karlis D., Wets G., 2008. Studying the effect of weather conditions on daily crash counts using a discrete time-

443 series model. *Accid. Anal. Prev.* 40: p. 1180 – 1190.

444 Brunson, C., Fotheringham, A.S., Charlton, M., 1996. Geographically weighted regression: a method for exploring

445 spatial non-stationarity. *Geographical Analysis,* p. 281-298.

446 Carroll, P.S., 1971. *Techniques for the use of driving exposure information in highway safety research.* HSRI, University

447 of Michigan.

448 Chapman, R.A., 1973. The concept of exposure. *Accid. Anal. Prev.* 5 (2), p. 95-110.

449 Charlton, M., Fotheringham, A.S., 2009. *Geographically Weighted Regression: White Paper.* National Centre for

450 Geocomputation, National University of Ireland Maynooth.

451 DATASUS, 2014. Ministério da saúde - Departamento de informática do SUS - Sistema de Informações sobre

452 Mortalidade (2010) - Estatísticas Vitais. Available in: <tab-net.datasus.gov.br>. Accessed in: 15 abr. 2014.

453 de Guevara, F.L., Washington, S.P., Oh, J., 2004. Forecasting crashes at the planning level: simultaneous negative

454 binomial crash model applied in Tucson, Arizona. In: *Transportation Research Record, Journal of the Transportation*

455 *Research Board,* No. 1897. Transportation Research Board of the National Academies. Washington, D.C, p. 191-199.

456 Delen, D., Sharada, R Bessonov, M., 2006. Identifying Significant Predictors of Injury Severity in Traffic Accidents

457 Using a Series of Artificial Neural Networks. *Accid. Anal. Prev.* 38: p. 434-444.

458 DENATRAN. Departamento Nacional de Tránsito. Available in: <<http://www.denatran.gov.br/>>.

459 Doherty, S.T., Andrey, J.C., MacGregor, C., 1998. The situational risks of young drivers: the influence of passenger,

460 time of day and day of week on accident rates. *Accid. Anal. Prev.* 30 (1): p. 45-52.

461 Elvik, R., 2007. State of the art approaches to road accident black spot management and safety analysis of road networks.

462 Institute of Transport Economics Norwegian Centre for Transport Research – rapport 883, Oslo.

463 Findley, D., Zegeer, C., Sundstrom, C., Hummer, J., Rasdorf, W., 2012. Applying the Highway Safety Manual to two-

464 lane road curves. *J. Transp. Res. Forum* 51 (3), 25-38.

465 Fotheringham, A.S., Brunson, C., Charlton, M., 1996. The geography of parameter space: an investigation of spatial

466 non-stationarity. *International Journal of Geographical Information Systems.* 10: p. 605-627.

467 Fotheringham, A.S., Charlton, M., Brunson, C., 1997. Two techniques for exploring nonstationarity in geographical

468 data, *Geographical Systems.* 4: p. 59-82.

469 Fotheringham, A.S., Brunson, C., Charlton, M.E., 2002. *Geographically Weighted Regression: The Analysis of*

470 *Spatially Varying Relationship.* Wiley: New York.

471 Fristrøm, L., Ifver, J., Ingebrigtsen, S., Kulmala, R., Thomsen, L.K., 1995. Measuring the contribution of randomness,

472 exposure, weather, and daylight to the variation in road accident counts. *Accid. Anal. Prev.* 27: p. 1-20.

473 Golob, T.F., Recker, W.W., 2003. Relationship among urban freeway accidents, traffic flow, weather, and lighting

474 conditions. *J. Transp. Eng.* 129: p. 342-353.

475 Guo, L., Ma, Z., Zhang, L., 2008. Comparison of bandwidth selection in application of geographically weighted

476 regression: a case study. *Canadian Journal of Forest Research.* 38: p. 2526 – 2534.

477 Hadayeghi, A., Shalaby, A.S., Persaud, B.N., 2003. Macrolevel accident prediction models for evaluating safety of urban

478 transportation systems. *Transportation Research Record: Journal of the Transportation Research Board.* 1840: p. 87-

479 95.

480 Hadayeghi, A., Shalaby, A.S., Persaud, B.N., 2010. Development of planning level transportation safety tools using
481 Geographically Weighted Poisson Regression. *Accid. Anal. Prev.* 42: p. 676 – 688.

482 Hao, W., Kamga, C., Wan, D., 2016. The effect of time of day on driver's injury severity at highway-rail grade crossings
483 in the United States. *Journal of Traffic and Transportation Engineering*. 3 (I): p. 37-50.

484 Harvey, A.C., Durbin, J., 1986. The effects of seat belt legislation on british road casualties: a case study in structural
485 time series modelling. *Journal of the royal statistical society. Series A (general)*, 149: p. 187-227.

486 Hauer, E., 1982. Traffic conflicts and exposure. *Accid. Anal. Prev.* 14 (5), p. 359–364.

487 Hauer, E., 1995. On exposure and accident rate. *Traffic Eng. Contr.* 36 (3), p. 134–138.

488 Hauer, E., Ng, J.C.N., Lovell, J., 1996. Estimation of safety at signalized intersection. *Transport. Res. Board* 1285: p.
489 42–51.

490 IBGE, 2014. Instituto Brasileiro de Geografia e Estatística. Censo Demográfico, 2010. Available in:
491 <www.censo2010.ibge.gov.br/resultados>. Accessed in: 17 jul. 2014.

492 Jalayer, M., Zhou, H., 2013. A sensitivity analysis of crash prediction models input in the Highway Safety Manual. Paper
493 presented at the 2013 ITE Midwestern District Meeting, Milwaukee, WI.

494 Jovanis, P., Chang, H.L., 1986. Modeling the relationship of accidents to miles traveled. *Transportation Res. Rec.* 1068,
495 p. 42–51.

496 Karlaftis, M. G., Golias, I., 2002. Effects of road geometry and traffic volumes on rural roadway accident rates. *Accid.*
497 *Anal. Prev.* 34: p. 357-365.

498 Kloeden, C. N., Ponte, G., McLean A. J., 2001. Travelling speed and the risk of crash involvement on rural roads.
499 Department of Transport and Regional Services Australian Transport Safety Bureau, Report no. CR 204, ISSN 1445-
500 4467.

501 Kononov, J., Janson, B., 2002. Diagnostic Methodology for the detection of safety problems at intersections, *Proceedings*
502 *of the Transportation Research Board*, Washington D.C.

503 Kutner, M.H., Nachtsheim, C.J., Neter, J., 2004. *Applied Linear Regression Models*, 4th ed. McGraw-Hill.

504 Langley, J., Mullin, B., Jackson, R. Norton, R., 2000. Motorcycle engine size and risk of moderate to fatal injury from a
505 motorcycle crash. *Accid. Anal. Prev.* 32: p. 659-663.

506 Li, Z., Wang, W., Liu, P., Bigham, J.M. Ragland, D.R., 2013. Using Geographically Weighted Poisson Regression for
507 county-level crash modeling in California. *Safety Science*. 58: p. 89-97.

508 Martin, J.L., 2002. Relationship between crash rate and hourly traffic flow on interurban motorways. *Accid. Anal. Prev.*
509 34: p. 619-629.

510 Miaou, S.P., 1994. The relationship between truck accidents and geometric design of road sections: poisson versus
511 negative binomial regressions. *Accid. Anal. Prev.* 26: p. 471-482.

512 Miaou, S.P., Song, J.J., Mallick, B.K., 2003. Roadway traffic crash mapping: a space–time modeling approach. *J.*
513 *Transportation Stat.* 6 (1), p. 33–57.

514 Mitra, S., Washington, S., 2012. On the significance of omitted variables in intersection crash modeling. *Accid. Anal.*
515 *Prev.* 49: p. 439-448.

516 MOBIEL VLAANDEREN. Mobiliteit en openbare werken. Available in: <<http://www.mobielvlaanderen.be>>.

517 Movig, K. L. L., Mathijssen, M.P.M., Nagel, P.H.A., Van Egmond, T., de Gier, J.J., Leufkens, H.G.M., Egberts, A.C.G.,
518 2004. Psychoactive substance use and the risk of motor vehicle accidents. *Accid. Anal. Prev.* 36: p. 631-636.

519 Nakaya, T., Fotheringham, A.S., Brunsdon, C., Charlton, M., 2005. Geographically weighted Poisson regression for
520 disease association mapping. *Statistics in Medicine*. 24: p. 2695–2717.

521 Nilsson, G., 2004. Traffic safety dimensions and the power model to describe the effect of speed on safety. Phd, lund
522 institute of technology.

523 Odgen, K.W., 1996. *Safer roads: a guide to road safety engineering*, burlington, ashgate publishing company.

524 OpenStreetMap. Available in <<https://www.openstreetmap.org>>.

525 Pei, X., Wong, S.C., Sze, N.N., 2012. The roles of exposure and speed in road safety analysis. *Accid. Anal. Prev.* 48: p.
526 464-471.

527 Petridou, E., Moustaki, M., 2000. Human factors in the causation of road traffic crashes. *European Journal of*
528 *Epidemiology*. 16: p. 819-826.

529 Pirdavani, A., Brijs, T., Bellemans, T., Kochan, B., Wets, G., 2013a. Evaluating the Road Safety Effects of a Fuel Cost
530 Increase Measure by means of Zonal Crash Prediction Modeling. In: *Accid. Anal. Prev.* 50, p. 186-195.

531 Pirdavani, A., Brijs, T., Bellemans, T., Wets, G., 2013b. Spatial analysis of fatal and injury crashes in Flanders, Belgium:
532 Application of geographically weighted regression technique. In: The 92th Annual Meeting of Transportation
533 ResearchBoard, Washington, DC.

534 Pirdavani, A., Bellemans, T., Brijs, T., Wets, G., 2014. Application of geographically weighted regression technique in
535 spatial analysis of fatal and injury crashes. *Journal of Transportation Engineering*.

536 Pirdavani, A., Daniels, S., van Vlieden, K., Brijs, K., Kochan, B., 2016. Socioeconomic and sociodemographic
537 inequalities and their association with road traffic injuries. In: *Journal of Transport & Health*, 4, p. 152-161.

538 Qin, X., Ivan, J. N., Ravishanker, N., 2004. Selecting exposure measures in crash rate prediction for two-lane highway
539 segments. *Accid. Anal. Prev.* 36: p. 183-191.

540 Qin, X., Ivan, J.N., Ravishanker, N., 2006. Bayesian estimation of hourly exposure functions by crash type and time of
541 day. *Accid. Anal. Prev.* 38 (6): p. 1071 – 1080.

542 Redelmeier, D.A., Tibshirani, R.J., 1997. Association between cellular-telephone calls and motor vehicle collisions. *New
543 england journal of medicine*. 336: p. 453-458.

544 Rengarasu, T. M., Hagiwara, T., Hirasawa, M., 2007. Effects of road geometry and season on head-on and single-vehicle
545 collisions on rural two lane roads in hokkaido, japan. *Journal of the Eastern Asia Society for Transportation Studies*,
546 Vol. 7, p. 2860-2872.

547 Richter, M., Pape, H.C., Otte, D., Krettek, C., 2005. Improvements in passive car safety led to decreased injury severity
548 – a comparison between the 1970s and 1990s. *Injury*. 36: p. 484-488.

549 Robertson, L.S., 1996. Reducing death on the road: the effects of minimum safety standards, publicized crash tests, seat
550 belts, and alcohol. *American journal of public health*. 86: p. 31-34.

551 Saha, D.; Alluri, P.; Gan, A., 2015. Prioritizing Highway Safety Manual’s crash prediction variables using boosted
552 regression trees. *Accid. Anal. Prev.* 79: p. 133–144. Shankar, V., Mannering, F. Barfield, W., 1995. Effect of roadway
553 geometrics and environmental factors on rural freeway accident frequencies. *Accid. Anal. Prev.* 27: p. 371-389.

554 Shankar, V., Mannering, F. Barfield, W., 1995. Effect of roadway geometrics and environmental factors on rural freeway
555 accident frequencies. *Accid. Anal. Prev.* 27: p. 371-389.

556 Shariat-Mohaymany, A., Shahri, M., Mirbagheri, B., Matkan, A.A., 2015. Exploringspatial non-stationarity and varying
557 relationships between crash data andrelated factors using geographically weighted Poisson regression. *Trans. GIS* 19
558 (2), 321–337.

559 Stewart, D.E., 1998. Methodological approaches for the estimation evaluation, interpretation and accuracy assessment of
560 road travel ‘basic risk’, ‘relative risk’, and ‘relative risk odds-ratio’ performance measure. Indicators: a ‘risk analysis
561 and evaluation system model’ for measuring, monitoring, comparing and evaluating the level(s) of safety on Canada’s
562 roads and highways. Transport Canada Publication No. TP 13238 E.

563 Taylor, M. C., Lynam, D.A., Baruya, A., 2000. The effects of drivers’ speed on the frequency of road accidents. *Transport
564 Research Laboratory Report* 421, ISSN 0968-4107.

565 Tobler, W., 1970. A computer movie simulating urban growth in the Detroit region. *Economic Geography*. 46: p. 234–
566 40.

567 Valent, F., Schiava, F., Savonitto, C., Gallo, T., Brusafferro, S. Barbone, F. 2002. Risk factors for fatal road traffic
568 accidents in udine, Italy. *Accid. Anal. Prev.* 34: p. 71-84.

569 Xu, C., Liu, P., Wang, W., Li, Z., 2012. Evaluation of the impacts of traffic states on crash risks on freeways. *Accid.
570 Anal. Prev.* 47: p. 162– 171.

571 Xu, P., Huang, H., Dong, N., Wong, SC., 2017. Revisiting crash spatial heterogeneity: A bayesian spatially varying
572 coefficients approach. *Accid. Anal. Prev.* 98: p. 330-337.

573 Yau, K.K.W., 2004. Risk factors affecting the severity of single vehicle traffic accidents in hong kong. *Accid. Anal. Prev.*
574 36: p. 333-340.

575 Wang, C., Quddus, M.A., Ison, S.G., 2009. Impact of traffic congestion on road accidents: a spatial analysis of the M25
576 motorway in England. *Accid. Anal. Prev.* 41 (4): p. 798–808.

577 Washington, S., Karlaftis, M., Mannering, F., 2010. *Statistical and Econometric Methods for Transportation Data
578 Analysis*, 2nd ed. Chapman & Hall, Boca Raton.

579 Williamson, M., Jalayer, M., Zhou, H., Pour-Rouholamin, M., 2015. A Sensitivity Analysis of Crash Modification
580 Factors of Access Management Techniques in Highway Safety Manual. *Access Management Theories and Practices*,
581 76.

582 WHO, 2015. World Health Organization. Global Status Report on Road Safety 2015. In:
583 <http://www.who.int/violence_injury_prevention/road_safety_status/2015/en/>, accessed in September 2017.