

Prognostic Value and Reproducibility of Simplified Noninvasive  
Pulmonary Artery Pressure-Flow Relationships in Exercise Echocardiography

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1 **Prognostic Value and Reproducibility of Simplified Non-Invasive Pulmonary**  
2 **Artery Pressure-Flow Relationships in Exercise Echocardiography**

3

4 **Data availability**

5 The data underlying this article are available in the article. The anonymized data will be  
6 shared on reasonable request to the corresponding author

7

8 **WORDS: 2612/3500 (excluding references, figure legends and tables)**

9 **ABBREVIATIONS**

10	exPHT	Exercise Pulmonary Hypertension
11	mPAP	Mean Pulmonary Artery Pressure
12	CO	Cardiac Output
13	CPET	Cardiopulmonary Exercise Test
14	CPETecho	Cardiopulmonary Exercise Test combined with Exercise
15		Echocardiography
16	ROC	Receiver Operating Characteristic
17	AUC	Area Under the Curve
18	ICC	Intraclass Correlation Coefficient
19		

20 **ABSTRACT**

21 ***Background***

22 The non-invasive mean pulmonary artery pressure (mPAP) to cardiac output (CO) relationship  
23 has been associated with cardiovascular outcomes across various populations. This study  
24 compared different methodologies for assessing this non-invasive relationship during exercise,  
25 with a dual focus on reproducibility and prognostic value. Specifically, we sought to determine  
26 whether simplified one-point mPAP/CO measurements offer equal reproducibility without  
27 compromising prognostic performance in patients with unexplained exertional dyspnea  
28 undergoing exercise echocardiography.

29 ***Methods and Results***

30 This was a secondary analysis of a prospective, multicenter cohort study including 1,619  
31 patients referred for unexplained dyspnea to six dedicated clinics in Belgium between January  
32 2016 and March 2023. All patients underwent symptom-limited cardiopulmonary exercise  
33 testing with concurrent exercise echocardiography. The mPAP-CO relationship was assessed  
34 using four approaches: (1) three-point linear regression, (2) two-point slope (rest to peak), (3)  
35 one-point mPAP/CO ratio at low-level exercise, and (4) one-point mPAP/CO ratio at peak  
36 exercise. The primary outcome was a composite of all-cause mortality or heart failure  
37 hospitalization. Median follow-up was 28.2 months (IQR: 15.9-45.1). Prognostic performance  
38 was evaluated by comparing discrimination across methods using receiver operating  
39 characteristic (ROC) curves. Reproducibility was assessed by intra-observer and inter-observer  
40 agreement.

41 The one-point at peak exercise showed the highest discriminative performance (AUC 0.715),  
42 with modest but significant improvement over the two-point (AUC 0.685,  $p = 0.019$ ) and three-  
43 point methods (AUC 0.686,  $p = 0.028$ ). The one-point at low-level exercise showed similar  
44 accuracy (AUC 0.713). Reproducibility was highest for the one-point ratios, with narrower  
45 intra- and inter-observer limits of agreement ( $p < 0.001$ ) compared to slope-based approaches.

46 ***Conclusions***

47 Simple, one-point mPAP/CO ratios measured during exercise offer prognostic performance  
48 comparable to slope-based methods, with higher reproducibility. This supports adoption as a  
49 practical and robust approach for evaluating exercise pulmonary hypertension.

50

51

52	<b>Keywords:</b>
53	Unexplained dyspnea
54	Exercise echocardiography
55	Exercise pulmonary hypertension
56	Risk Assessment
57	
58	

59 **Clinical Relevance to Patient Care:**

- 60 • Simplified mPAP/CO ratios offer a practical, non-invasive tool to assess exercise  
61 pulmonary hypertension
- 62 • Low-level exercise ratios provide a reliable alternative when peak measurements are  
63 not feasible
- 64 • High reproducibility supports consistent use across clinicians and institutions  
65

## 66 INTRODUCTION

67 Exertional dyspnea is a prevalent yet diagnostically challenging symptom encountered across  
68 multiple medical specialties.<sup>1-3</sup> Despite the absence of abnormal resting findings, the symptom  
69 is frequently associated with exercise-induced cardiac or pulmonary pathology and has a well-  
70 documented prognostic significance.<sup>4,5</sup> Since resting evaluations are commonly inconclusive,  
71 exercise-based assessments have become central to diagnostic work-up in this setting and  
72 dedicated dyspnea clinics have emerged worldwide.<sup>6,7</sup>

73 Within this diagnostic landscape, exercise pulmonary hypertension (exPHT) represents a  
74 clinically important entity, affecting approximately 41% of patients with dyspnea, occurring as  
75 either a primary cause or secondary to left heart disease, pulmonary disease, or systemic  
76 conditions.<sup>4</sup> Historically, exPHT was defined as a mean pulmonary artery pressure (mPAP) > 30  
77 mmHg during exercise. However, this threshold was later abandoned, as healthy individuals  
78 can exceed it at high cardiac outputs (CO).<sup>8</sup> Accordingly, interpretation of pressure-flow  
79 relationships, rather than absolute pressure alone, has become essential to distinguish  
80 physiological from pathological responses to exercise.<sup>9,10</sup> Current pulmonary hypertension  
81 (PHT) guidelines define exPHT invasively as an abnormal mPAP-to-CO slope >3.0 mmHg/L/min  
82 during exercise.<sup>11</sup>

83 ExPHT independently predicts adverse outcomes in patients with exertional dyspnea.<sup>4,12,13</sup>  
84 Although invasive hemodynamic measurements remain the gold standard, previous studies  
85 have demonstrated the prognostic value of non-invasively assessed exPHT using exercise  
86 echocardiography across various cardiac conditions on top of resting evaluations.<sup>12,14,15</sup>

87 Beyond prognostic risk stratification, pulmonary pressure–flow responses during exercise may  
88 also reflect underlying pathophysiological mechanisms. Recent work has linked abnormal  
89 mPAP/CO slopes to impaired ventricular-arterial interactions and systemic factors such as iron  
90 deficiency, irrespective of LVEF phenotype.<sup>16,17</sup>

91 Despite its clinical importance, there is currently no standardized non-invasive method to  
92 quantify the mPAP-CO relationship during exercise. The reproducibility and prognostic  
93 implications of different non-invasive approaches remain largely unexplored. To date, no large  
94 study has directly compared multiple non-invasive methods for assessing mPAP/CO in a well-  
95 characterized clinical population.

96 In this study, we aimed to compare four commonly used non-invasive methods for calculating  
97 the mPAP-CO relationship in patients with unexplained exertional dyspnea undergoing  
98 exercise echocardiography. We hypothesized that a simplified, one-point mPAP/CO ratio  
99 during exercise would improve reproducibility without compromising prognostic value.

100

## 101 METHODS

### 102 *Study design and population*

103 This was a secondary analysis of a multicenter cohort study conducted at six multidisciplinary  
104 dyspnea clinics in Belgium. It included patients with unexplained dyspnea referred between  
105 January 2016 and March 2023 for cardiopulmonary exercise testing with concurrent exercise  
106 echocardiography (CPETecho). Clinical procedures were standardized across all six centers, as  
107 previously described.<sup>7,18</sup> Data was systematically collected during routine care. Patients with a  
108 clear resting diagnosis were excluded (**Figure 1**). Only patients with complete mPAP and CO  
109 measurements at rest, low-level, and peak exercise were included in the analysis.

110 The primary composite outcome was the time to first occurrence of all-cause mortality or  
111 heart failure hospitalization. Outcome was assessed via retrospective chart review of all  
112 available medical records. Heart failure hospitalization was defined as an unplanned admission  
113 requiring intravenous diuretics, inotropes, or vasodilators. Patients who were lost to follow-up  
114 within the first year and did not experience an event were excluded to ensure complete  
115 outcome classification.

116 The study protocol received approval from local ethics committees. Due to the retrospective  
117 nature of the analysis, the requirement for written informed consent was waived. This study  
118 complies with the principles of the Declaration of Helsinki and Good Clinical Practice  
119 guidelines.

#### 120 ***Assessment of exercise pulmonary pressures and cardiac output***

121 All patients underwent a maximal symptom-limited CPETecho on a semi-supine bicycle  
122 positioned at a 45° angle, using individualized ramp protocols (5-20 W/min) targeting a total  
123 duration of 10-15 minutes. Pulmonary pressures and CO were assessed at three standardized  
124 stages: rest, low-intensity exercise (dedicated hold stage), and peak exercise, as previously  
125 described.<sup>18</sup>

126 The low-intensity stage was defined as the period after surpassing the first ventilatory  
127 threshold but before reaching a heart rate (HR) of 100 bpm. Peak exercise measurements were  
128 obtained during a brief hold stage immediately before exhaustion, typically defined as a  
129 respiratory exchange ratio (RER) > 1.05, or earlier in case of symptom-limitation.

130 Stroke volume (SV) was calculated as the product of the time-velocity integral (VTI) at the left  
131 ventricular outflow tract and its cross-sectional area. CO was derived as SV x HR. Systolic  
132 pulmonary artery pressure (sPAP) was estimated from the maximal tricuspid regurgitation  
133 velocity (TRV) using continuous wave Doppler:

$$134 \text{ sPAP} = 4 \times \text{TRV}^2$$

135 Mean pulmonary artery pressure (mPAP) was calculated from sPAP using the Chemla formula  
136 (without an estimate of right atrial pressure):

$$137 \text{ mPAP} = 0.61 \times \text{sPAP} + 2 \text{ mmHg.}^{19}$$

138 To enhance the TR envelope, an agitated colloid bolus (1-3 mL, tailored to TR quality) was  
139 routinely administered at rest and during exercise.<sup>20</sup> This method has been previously  
140 validated against invasive exercise hemodynamics.<sup>21</sup> All measurements were averaged over at  
141 least three cardiac cycles and recorded in a standardized data sheet. Image acquisition was

142 performed using a Vivid E9 or E95 ultrasound system (GE Vingmed Ultrasound AS), and offline  
143 analysis was performed using *EchoPAC* (v112, GE Healthcare).

#### 144 ***Evaluation of the mPAP-CO relationship***

145 The pressure-flow relationship between the pulmonary vasculature and CO was evaluated  
146 using four different non-invasive methods:

- 147 **1. Three-point slope** (linear regression across rest, low and peak),
- 148 **2. Two-point slope** (rest to peak),
- 149 **3. One-point mPAP/CO ratio** at low exercise,
- 150 **4. One-point mPAP/CO ratio** at peak exercise.

151 The three-point slope was derived from linear regression with mPAP as the dependent variable  
152 (y) and CO as the independent variable (x). Calculation was implemented via an embedded  
153 Excel® formula within a report-generating database. The two-point slope was calculated as the  
154 change in mPAP divided by the corresponding change in CO between rest and peak exercise.

#### 155 ***Inter- and intra-observer variability analysis***

156 To evaluate reproducibility, 40 CPETecho exams from X Hospital were randomly selected. Three  
157 experienced observers from different centers (X.X – X, Y.Y – Y, Z.Z – Z) independently measured  
158 mPAP and CO. All observers were blinded to prior measurements and to post-hoc slope  
159 calculations. To assess intra-observer variability, each observer repeated their measurements  
160 after a three-week interval.

#### 161 ***Statistical Analysis***

162 Prognostic accuracy of the four mPAP/CO methods was assessed using receiver operating  
163 characteristic (ROC) curves. Areas under the curve (AUCs) were compared using DeLong’s test.  
164 Sensitivity, specificity, and optimal cut-off values were identified using the Youden index.

165 Reproducibility was evaluated using intraclass correlation coefficients (ICC) with 95%  
166 confidence intervals (CIs) for inter- and intra-observer agreement. Fleiss’ kappa was used for  
167 agreement among all raters, and Cohen’s kappa for intra-rater agreement, considering the  
168 identification of exPHT (mPAP/CO > 3.0 mmHg/L/min). Agreement between paired  
169 measurements was further assessed using Bland-Altman plots, reporting bias, and limits of  
170 agreement ( $\pm 1.96$  SD). Method comparisons were evaluated by analyzing differences in mean  
171 bias using one-way ANOVA and unpaired t-tests, while differences in variability among  
172 methods were assessed using Brown-Forsythe, Bartlett’s, and F-tests.

173 Cox proportional hazards regression models were used to evaluate the association between  
174 the four methods and the composite endpoint of all-cause death or hospitalization for heart  
175 failure. Univariable models were first performed, followed by multivariable models adjusting  
176 for age and VO<sub>2</sub>peak (as a continuous variable) (model B) and predicted VO<sub>2</sub>peak (categorical  
177 – with 80% as the cut-off for reduced exercise capacity) and RER (categorical – with 1.05 as the  
178 cut-off for maximal effort) (model C). Results are reported as hazard ratios (HR) with  
179 corresponding 95% confidence intervals (CI). Logistic regression was performed, including  
180 each method, resting mPAP, and their interaction term.

181 All statistical analyses were conducted using *SPSS Statistics* (v29, IBM, Chicago, IL) and *MedCalc*  
182 *Software*. Figures were generated with *GraphPad Prism* (v10, GraphPad Software, San Diego,  
183 CA). P-values < 0.05 were considered statistically significant.

## 184 **RESULTS**

### 185 ***Study population***

186 Eleven percent (521 patients) of the initial cohort were excluded due to non-feasibility or  
187 incomplete pulmonary pressure or CO measurements at one or more exercise levels (**Figure**  
188 **1**). The final study cohort consisted of 1,619 patients with unexplained dyspnea. Baseline  
189 characteristics are summarized in **Table 1**. The mean age was 62 ± 15 years and 54% were  
190 women. The median NT-proBNP was 120 [52 - 209] ng/L.

191 Over a median follow-up of 28 months (IQR 16-45), 137 patients (8.5%) experienced the  
192 primary composite outcome, which included 66 deaths (4.1%) and 84 heart failure  
193 hospitalizations (5.2%).

### 194 ***Prognostic value of mPAP/CO methods***

195 Among the four methods evaluated, the one-point mPAP/CO ratio at peak exercise  
196 demonstrated the highest prognostic performance, with an AUC of 0.715 (95% CI: 0.692–  
197 0.737). This was modest but significantly higher than the 2-point slope from rest to peak (AUC:  
198 0.685; 95%CI 0.662–0.708; p=0.028) and the 3-point regression (AUC: 0.686, 95% CI: 0.663–  
199 0.709; p=0.028) (**Table 2, Figure 2**).

200 The one-point mPAP/CO ratio at low-level exercise showed comparable prognostic  
201 performance (AUC: 0.713; 95% CI: 0.690–0.735) to that observed at peak exercise. The cut-off  
202 values ranged from 3.4 to 3.6 mmHg/L/min. Sensitivity, specificity, positive predictive value  
203 (PPV), negative predictive value (NPV), and optimal cut-off values for each method are  
204 summarized in **Table 2**.

### 205 ***Reproducibility of mPAP/CO Methods***

206 One-point mPAP/CO ratio assessments at low and peak exercise demonstrated excellent intra-  
207 observer reliability, with ICCs of 0.903 (95% CI: 0.863–0.932) and 0.923 (95% CI: 0.890–0.946),  
208 respectively (**Table 3** and **Supplementary Figure 1**). The 2-point slope and 3-point regression  
209 methods showed good intra-observer reliability, with ICCs of 0.874 (95% CI: 0.822–0.911) and  
210 0.857 (95% CI: 0.799–0.899).

211 Inter-observer reliability was also good across all methods, with ICCs ranging from 0.802 to  
212 0.880 (**Table 4** and **Supplementary Figure 1**). The highest Fleiss' kappa was observed for the  
213 peak one-point mPAP/CO ratio method (0.825, 95% CI: 0.610–1.000) (**Table 4** and  
214 **Supplementary Figure 2**).

215 Bland-Altman plots for intra-observer agreement (**Figure 3A**) demonstrated similar bias  
216 between methods (p=0.43), but revealed significantly narrower limits of agreement for the  
217 one-point approaches compared to the 2-point and 3-point slope methods (p<0.001),  
218 indicating superior repeatability. Similar patterns were observed in inter-observer analysis

219 **(Figure 3B)**, where slope-based methods exhibited broader limits of agreement ( $p < 0.001$ ),  
220 reflecting reduced reproducibility.

### 221 ***Clinical Profile by one-point mPAP/CO ratio***

222 Patients with a one-point mPAP/CO ratio at peak of  $>3.6$  mmHg/L/min were older and more  
223 frequently had hypertension, diabetes, and atrial fibrillation. These patients also had larger  
224 left atrial volumes, higher NT-proBNP concentrations, and more advanced diastolic dysfunction  
225 at rest (**Table 1**).

### 226 ***Sensitivity analyses***

227 Cox proportional hazards regression was performed as a sensitivity analysis to account for the  
228 time-to-event nature of the data and to assess the prognostic value of the four different  
229 methods (**Supplementary Table 1**). All the methods were predictors of the composite endpoint  
230 in a univariable analysis (all  $p < 0.001$ ).

231 The one-point approach at low and peak exercise remained significant in a multivariable model  
232 A, which adjusted for age and cardiorespiratory fitness (CRF) as a continuous variable ( $p =$   
233  $0.003$  and  $0.04$ , respectively), whereas the 2-point and 3-point methods were not ( $p = 0.14$   
234 and  $0.06$ , respectively) (**Supplementary Table 1**). When CRF was modeled categorically  
235 ( $VO_{2peak} < 80\%$  vs.  $\geq 80\%$  of predicted) in model B, mPAP/CO remained a significant predictor  
236 across all methods (all  $p < 0.001$ ). In model C, which incorporated exercise intensity ( $RER <$   
237  $1.05$  vs.  $\geq 1.05$ ), mPAP/CO continued to demonstrate independent prognostic value across all  
238 models ( $p < 0.001$ ).

239 To assess whether a non-linear pressure-flow relationship influenced the prognostic value of  
240 the mPAP/CO, the association between CO and mPAP was evaluated across exercise stages. A  
241 linear relationship was observed between mPAP and the relative percentage of peak CO  
242 (**Supplementary Figure 3**). Resting mPAP had a small but significant effect on the prognostic  
243 performance of the mPAP-CO relationship, but this was similar across mPAP/CO methods  
244 (**Supplementary Table 2**).

245

## 246 **DISCUSSION**

247 Accurate assessment of exercise pulmonary hypertension (exPHT) is crucial for risk  
248 stratification and management of patients with unexplained dyspnea. In this large multicenter  
249 cohort study, we systematically compared the prognostic performance and reproducibility of  
250 four non-invasive methods for evaluating the mPAP-CO relationship during exercise. Our  
251 findings support two key conclusions: first, one-point mPAP/CO ratios, obtained either at low-  
252 level or peak exercise, provide prognostic performance comparable to more complex slope-  
253 based approaches; second, these simplified methods demonstrate superior reproducibility,  
254 with higher inter- and intra-rater reliability and narrower limits of agreement. Taken together,  
255 these findings support the use of one-point mPAP/CO ratio measurements as a practical,  
256 robust, and prognostically relevant strategy for evaluating exPHT in patients with unexplained  
257 dyspnea.

258 Recent studies have underscored the utility of non-invasive exercise testing for assessing  
259 pulmonary vascular response in patients with unexplained dyspnea and a range of  
260 cardiorespiratory disorders, complementing traditional resting evaluations that include  
261 pulmonary pressure assessment.<sup>12,14,15</sup> Non-invasive mPAP/CO slope estimates have been  
262 validated against invasive hemodynamics and shown to be feasible in up to 87% of cases when  
263 using agitated contrast.<sup>20,21</sup> However, no consensus exists on a standardized non-invasive  
264 approach, and calculation methods vary widely across studies.

265 This variability is clinically relevant: in an invasive study by Godinas et al., different criteria for  
266 defining exPHT led to only moderate concordance, with disagreement up to 22% of cases.<sup>22</sup>  
267 Such heterogeneity underscores the need for methodological consistency and supports the  
268 potential advantages of simplified, one-point approaches.

269 In our study, although the differences in AUC values between methods were modest (**Figure**  
270 **2**), one-point measures demonstrated consistently better reproducibility (**Figure 3**,  
271 **Supplementary Figures 1-2**). This advantage may reflect the vulnerability of slope-based  
272 methods to compounding error, as they rely on multiple measurements that each carry some  
273 degree of imprecision.<sup>23,24</sup> In contrast, one-point estimates are inherently less susceptible to  
274 cumulative variability. This is consistent with the data in **Supplementary Figure 4**, which did  
275 not identify any specific component of the mPAP-CO relationship, such as stroke volume, CO,  
276 or mPAP, as a primary contributor to the observed differences between methods.

277 Sensitivity analysis confirmed that the one-point mPAP/CO ratio at low and peak exercise  
278 remained independently associated with outcome, even after adjustment for age and  
279 cardiorespiratory fitness (**Supplementary Table 1**). All methods were independently  
280 associated with outcome at different levels of physical effort and fitness.

281 Notably, neither the curvilinear pressure–flow relationship previously described in other  
282 studies<sup>25,26</sup> (**Supplementary Figure 3**), nor a “higher take-off” pattern associated with resting  
283 mPAP (**Supplementary Table 2**), appeared to explain the observed differences between  
284 methods. These findings support the notion that technical and methodological factors, rather  
285 than physiological mechanisms, may underlie the superior performance of one-point  
286 measures.

287 Stratification by mPAP/CO ratio revealed a more adverse clinical profile in patients with a  
288 steeper pressure–flow response, consistent with an HFpEF phenotype (**Table 1**). We recently  
289 demonstrated that this prognostic association was independent of and incremental to NT-  
290 proBNP, HFpEF probability scores, diastolic stress testing, and absolute pulmonary pressures,  
291 supporting the notion that these individuals may represent an early or exertional HFpEF  
292 phenotype not captured by current non-invasive diagnostic criteria.<sup>15</sup> Notably, patients with  
293 an elevated exercise mPAP/CO ratio but negative HFpEF scores showed no evidence of  
294 pulmonary vascular disease: they had normal ventilatory efficiency and reserve, and no  
295 hypoxemia during exercise, further suggesting a predominantly cardiac origin of their  
296 abnormal hemodynamic response.

297 ***Clinical implications***

298 Our findings have direct implications for practice. Simplified one-point mPAP/CO ratio ratios  
299 offer a feasible and reproducible tool for evaluating total pulmonary resistance during exercise.  
300 By reducing methodological complexity and improving reproducibility, this approach facilitates  
301 broader implementation in both routine clinical follow-up and the evaluation of treatment  
302 effects in research.

303 Importantly, when peak exercise measures are not feasible, due to early termination, technical  
304 limitations, or patient discomfort, low-level one-point measurements provide comparable  
305 prognostic information and serve as a reliable alternative.

306 Higher inter- and intra-rater agreement further supports the consistency of one-point  
307 approaches across readers and institutions. For longitudinal research or clinical trials, their  
308 reproducibility makes them particularly suited for tracking change over time.

### 309 ***Study limitations***

310 The study has several limitations. First, outcome data was collected retrospectively through  
311 chart review, which may have resulted in incomplete event capture. However, we used a  
312 nationwide digital platform that consolidates medical records from all Belgian hospitals,  
313 minimizing the risk of missed outcomes.

314 Second, patients with incomplete hemodynamic data at all three stages were excluded,  
315 potentially introducing selection bias. Nonetheless, this allowed for consistent within-subject  
316 comparisons across methods and improved data integrity.

317 Third, the ROC analysis did not account for the time-to-event nature of the data. However, a  
318 time-to-event analysis using Cox proportional hazards regression was performed in our  
319 primary investigation of this cohort<sup>15</sup>, and a similar sensitivity analysis in the current study  
320 yielded consistent results, supporting the robustness of the prognostic comparisons across  
321 methods.

322 Fourth, there was an unbalanced distribution of patients across centers, with more than 80%  
323 recruited at a single site. This precluded robust center-adjusted analysis and highlights the  
324 need for external validation of the cut-off values in more balanced multicenter cohorts.

325 Lastly, all hemodynamic parameters were derived non-invasively. Although invasive measures  
326 remain the gold standard for diagnosing exPHT, the echocardiographic protocols used here  
327 have been previously validated against invasive exercise hemodynamics.<sup>21</sup> Moreover, the  
328 primary aim of this study was not to assess diagnostic accuracy, but to evaluate prognostic  
329 performance. All measurements were performed by experienced personnel using  
330 standardized protocols, and high reproducibility across observers supports the validity of our  
331 findings.

### 332 **CONCLUSION**

333 In this large, multicenter cohort of patients with unexplained exertional dyspnea, simple one-  
334 point mPAP/CO ratios during exercise provided prognostic value comparable to more complex  
335 two- or three-point slope-based methods, while offering significantly greater reproducibility.  
336 These findings support the use of the mPAP/CO ratio as a practical, reproducible, and

337 physiologically meaningful approach for the non-invasive evaluation of exercise pulmonary  
338 hypertension. Superior reproducibility is crucial for longitudinal follow-up and assessment of  
339 therapeutic interventions.

340

341

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425

426 **Figure captions**

427 **Figures (separate files)**

428

429 **Central Figure (Visual Take Home Graphic):** Study design and key findings

430 **Figure 1.** Flowchart of study population

431 **Figure 2.** Receiver operating characteristic (ROC) analysis comparing the prognostic accuracy  
432 of different methods for assessing the mPAP-CO relationship

433 **Figure 3.** Bland-Altman plots for intra- and inter-rater agreement

434

1,619 patients with unexplained dyspnea underwent CPETechno

- Four methods to assess mPAP/CO relationship
- 1-point ratio at Peak
  - 1-point ratio at Low
  - 2-point (Rest to Peak)
  - 3-point Regression

**Composite endpoint:** all-cause mortality or HFH

One-point mPAP/CO ratios are

*Simpler*

Excel® vs  $mPAP/CO = 30/10 = 3$

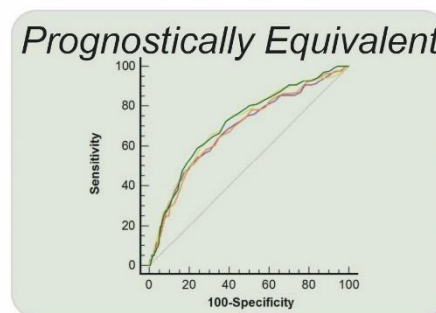
*More Reproducible*

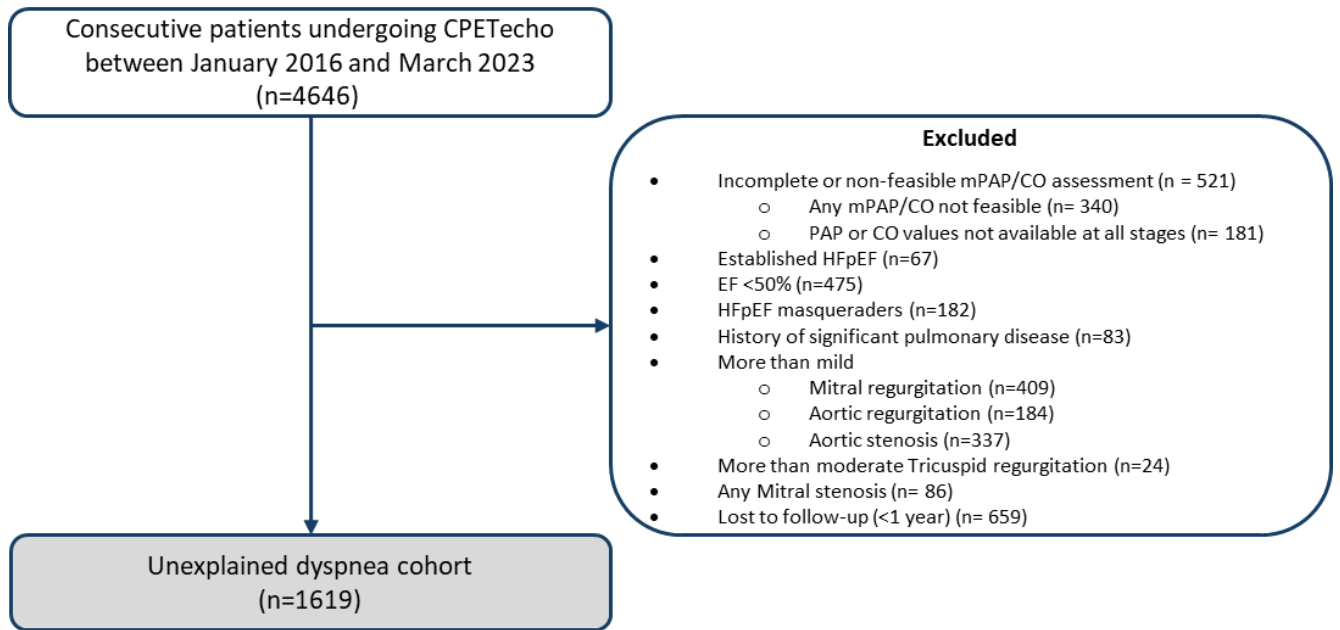
*Physiologically Sound*

Pulmonary Vascular      Left Atrium

RV → [ECG trace] → [ECG trace] → LV

Exercise Total Pulmonary Resistance

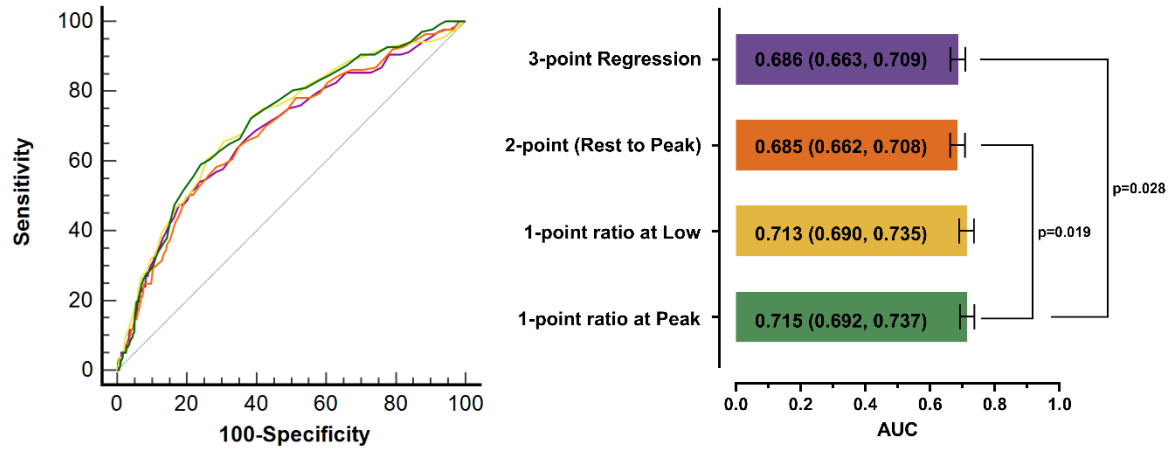




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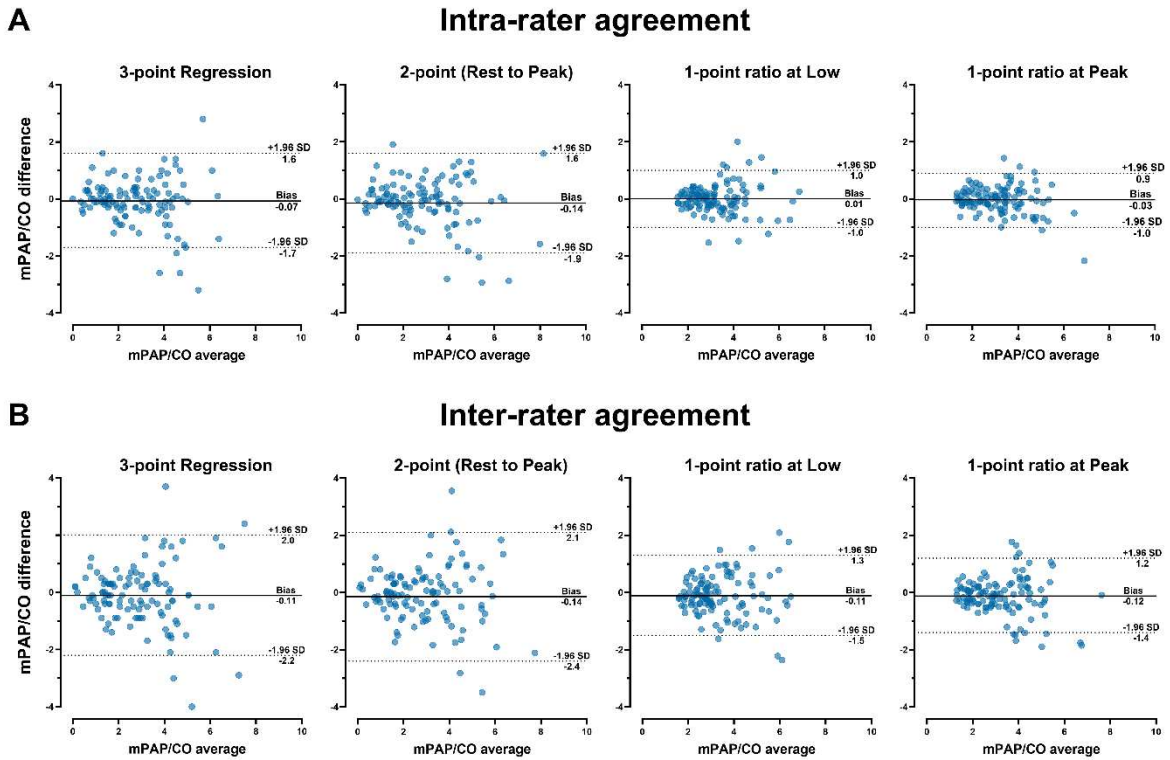
439 **Figure 1:** CPETecho, Cardiopulmonary Exercise Test combined with echocardiography; CO, cardiac output; PAP,   
 440 pulmonary artery pressure; EF, ejection fraction; HFpEF, Heart Failure with Preserved Ejection Fraction

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**Figure 2:** Receiver operating characteristic (ROC) curves for the different mPAP/CO methods. Bars indicate the area under the curve (AUC) and 95% confidence intervals. P-values reflect the statistical significance of differences in AUC between methods.



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447 **Figure 3:** Bland-Altman plots illustrating the mean difference (bias) and 95% limits of agreement ( $\pm 1.96$  SD) for  
 448 four different methods to assess the mPAP-CO relationship. Panel A shows intra-rater agreement; Panel B shows  
 449 inter-rater agreement. Each point represents an individual paired comparison between two measurements by the  
 450 same rater (Panel A) or between two different raters (Panel B).

451 CO, cardiac output; mPAP, mean pulmonary artery pressure; SD, standard deviation

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453 **Table captions**

454 **Tables (separate files)**

455

456 **Table 1.** Patient characteristics

457 **Table 2.** Diagnostic accuracy according of methods used to calculate the mPAP-CO  
458 relationship.

459 **Table 3.** Intra-rater reliability

460 **Table 4.** Inter-rater reliability

	Total cohort	mPAP/CO ratio $\leq 3.6$	mPAP/CO ratio $> 3.6$	p-value
Number of patients	1619	1159	460	
<b>Demographics</b>				
Age (years)	62 $\pm$ 15	58 $\pm$ 16	72 $\pm$ 9	<0.001
Female, n (%)	872 (54)	578 (49.9%)	164 (35.7%)	<0.001
Systolic blood pressure (mmHg)	137 $\pm$ 22	136 $\pm$ 24	140 $\pm$ 25	0.003
Diastolic blood pressure (mmHg)	79 $\pm$ 14	80 $\pm$ 14	77 $\pm$ 16	<0.001
Body Mass Index (kg.m <sup>-2</sup> )	27 $\pm$ 5	27 $\pm$ 5	27 $\pm$ 5	0.030
Hypertension, n (%)	728 (45)	441 (38)	287 (62)	<0.001
Diabetes, n (%)	189 (12)	121 (10)	68 (15)	0.016
Atrial Fibrillation history, n (%)	296 (18)	145 (13)	151 (33)	<0.001
<b>Biomarker</b>				
NT-proBNP (ng.L <sup>-1</sup> )	120 (52, 290)	84 (40,180)	290 (140,570)	<0.001
<b>Echocardiography (Rest)</b>				
LAVi (mL.m <sup>-2</sup> )	21 $\pm$ 14	20 $\pm$ 13	24 $\pm$ 17	<0.001
E/A	1.1 $\pm$ 0.6	1.1 $\pm$ 0.6	1.1 $\pm$ 0.7	0.081
E/e'	11.2 $\pm$ 4.6	10.0 $\pm$ 3.5	14.3 $\pm$ 5.9	<0.001
TRPG (mmHg)	22.7 $\pm$ 6.1	21.1 $\pm$ 4.8	26.7 $\pm$ 7.1	<0.001
<b>Outcome</b>				
Deaths, n (%)	66 (4.1)	32 (2.8%)	34 (7.4%)	<0.001
Heart failure hospitalizations, n (%)	84 (5.2)	29 (2.5%)	55 (12.0%)	<0.001
Composite outcome, n (%)	137 (8.5)	56 (4.8%)	81 (17.6%)	<0.001

461 **Table 1:** Baseline characteristics of the total cohort (n=1619), stratified by one-point mPAP/CO ratio at peak  
462 exercise ( $\leq 3.6$  vs.  $> 3.6$  mmHg/L/min). Data presented as mean  $\pm$  standard deviation, median (25<sup>th</sup> – 75<sup>th</sup>  
463 percentiles), or number (%). P-values reflect between-group comparisons using t-tests or chi-square tests as  
464 appropriate. NT-proBNP, N-terminal pro-B-type natriuretic peptide; LAVi, left atrium volume index; TRG,  
465 Tricuspid Regurgitation Pressure Gradient.

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	AUC (95%CI)	Optimal Cut-off	Sensitivity (%)	Specificity (%)	PPV (%)	NPV (%)
<b>3-point Regression</b>	0.686 (0.663-0.709)	> 3.5	54.0	76.4	17.5	94.7
<b>2-point (Rest to Peak)</b>	0.685 (0.662-0.708)	> 3.4	58.4	71.5	16.0	94.9
<b>1-point ratio at Low</b>	0.713 (0.690-0.735)	> 3.5	65.7	69.3	16.6	95.6
<b>1-point ratio at Peak</b>	0.715 (0.692-0.737)	> 3.6	59.1	76.1	18.7	95.2

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**Table 2:** Prognostic performance of the four methods used to assess the mPAP-CO relationship. Area under the curve (AUC) with 95% confidence intervals (CI), and optimal cut-off values based on the Youden Index are presented for each method. Sensitivity and specificity are reported according to the respective optimal cut-off values. Positive predictive value (PPV) and negative predictive value (NPV) were calculated assuming a prevalence of 8.5% (137 events among 1619 patients).

CI, confidence interval; CO, cardiac output; mPAP, mean pulmonary artery pressure.

mPAP/CO method	First	Second	R (Pearson)	ICC	Cohen's Kappa*
<b>3-point Regression</b>	2.7 ±1.5	2.8 ±1.6	0.859	0.857 (0.799-0.899)	0.875
<b>2-point (Rest to Peak)</b>	2.9 ± 1.7	3.0 ± 1.9	0.884	0.874 (0.822-0.911)	0.841
<b>1-point ratio at Low</b>	3.3 ± 1.2	3.2 ± 1.2	0.902	0.903 (0.863-0.932)	0.772
<b>1-point ratio at Peak</b>	3.1 ± 1.2	3.1 ± 1.3	0.926	0.923 (0.890-0.946)	0.842

483 **Table 3:** Intra-observer reliability for the four methods used to calculate the mPAP-CO relationship. Data is  
484 expressed as mean ± SD for the first and second measurements by the same rater. Pearson correlation coefficients  
485 (R), intraclass correlation coefficients (ICC) with 95% confidence intervals, and Cohen's kappa values are reported.

486 \* Categorization of mPAP/CO > 3.0 mmHg/L/min.

487 CO, cardiac output; ICC, intraclass correlation coefficient; mPAP, mean pulmonary artery pressure.

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mPAP/CO method	Rater 1	Rater 2	Rater 3	ICC	Fleiss' Kappa*
<b>3-point Regression</b>	2.6 ± 1.4	3.1 ± 2.1	2.8 ± 1.5	0.802 (0.682-0.885)	0.681 (0.497-0.864)
<b>2-point (Rest to Peak)</b>	2.8 ± 1.6	3.2 ± 2.1	3.0 ± 2.1	0.828 (0.728-0.899)	0.718 (0.534-0.901)
<b>1-point ratio at Low</b>	3.1 ± 1.2	3.8 ± 1.5	3.2 ± 1.2	0.851 (0.559-0.938)	0.684 (0.501-0.868)
<b>1-point ratio at Peak</b>	3.0 ± 1.2	3.5 ± 1.5	3.2 ± 1.3	0.880 (0.747-0.941)	0.825 (0.641-1.000)

490 **Table 4:** Inter-observer reliability for the four methods used to calculate the mPAP-CO relationship. Data is  
491 expressed as mean ± SD for each rater. Intraclass correlation coefficients (ICC) with 95% confidence intervals and  
492 Fleiss' kappa values are reported.

493 \* Categorization of mPAP/CO > 3.0 mmHg/L/min

494 CO, cardiac output; ICC, intraclass correlation coefficient; mPAP, mean pulmonary artery pressure.

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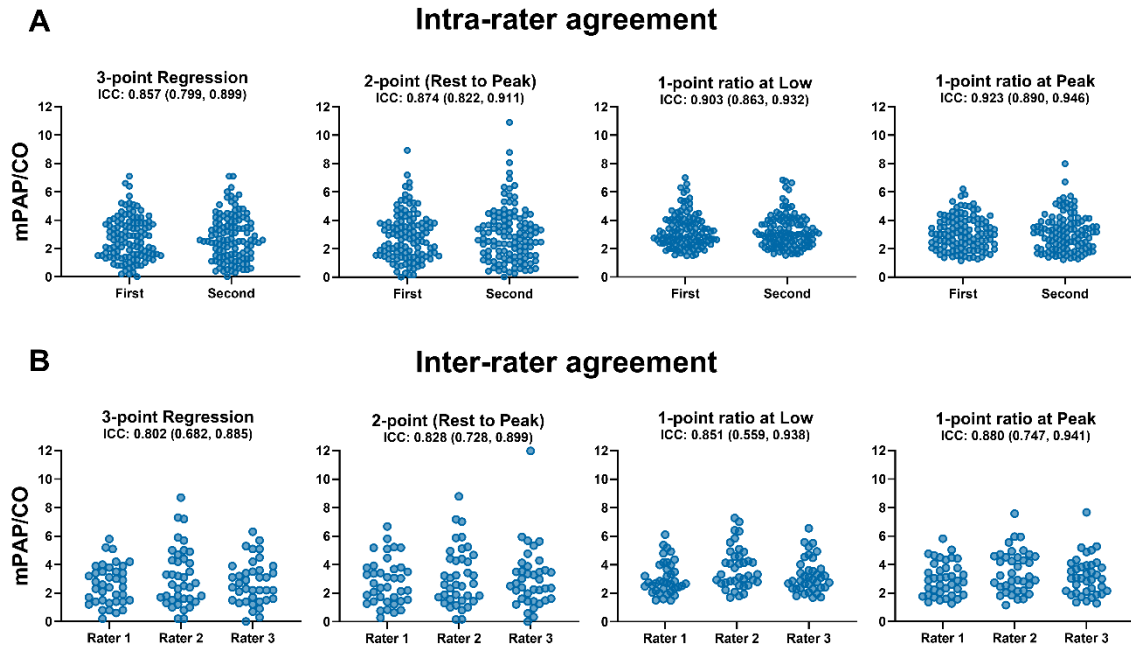
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504 **SUPPLEMENTARY MATERIALS**

505 **Supplementary Figure 1. Intra- and inter-rater variability by intraclass correlation coefficient**  
506 **(ICC)**



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508 *Supplementary Figure 1: Data expressed as intraclass correlation coefficient (ICC) with its corresponding 95%*  
509 *confidence interval.*

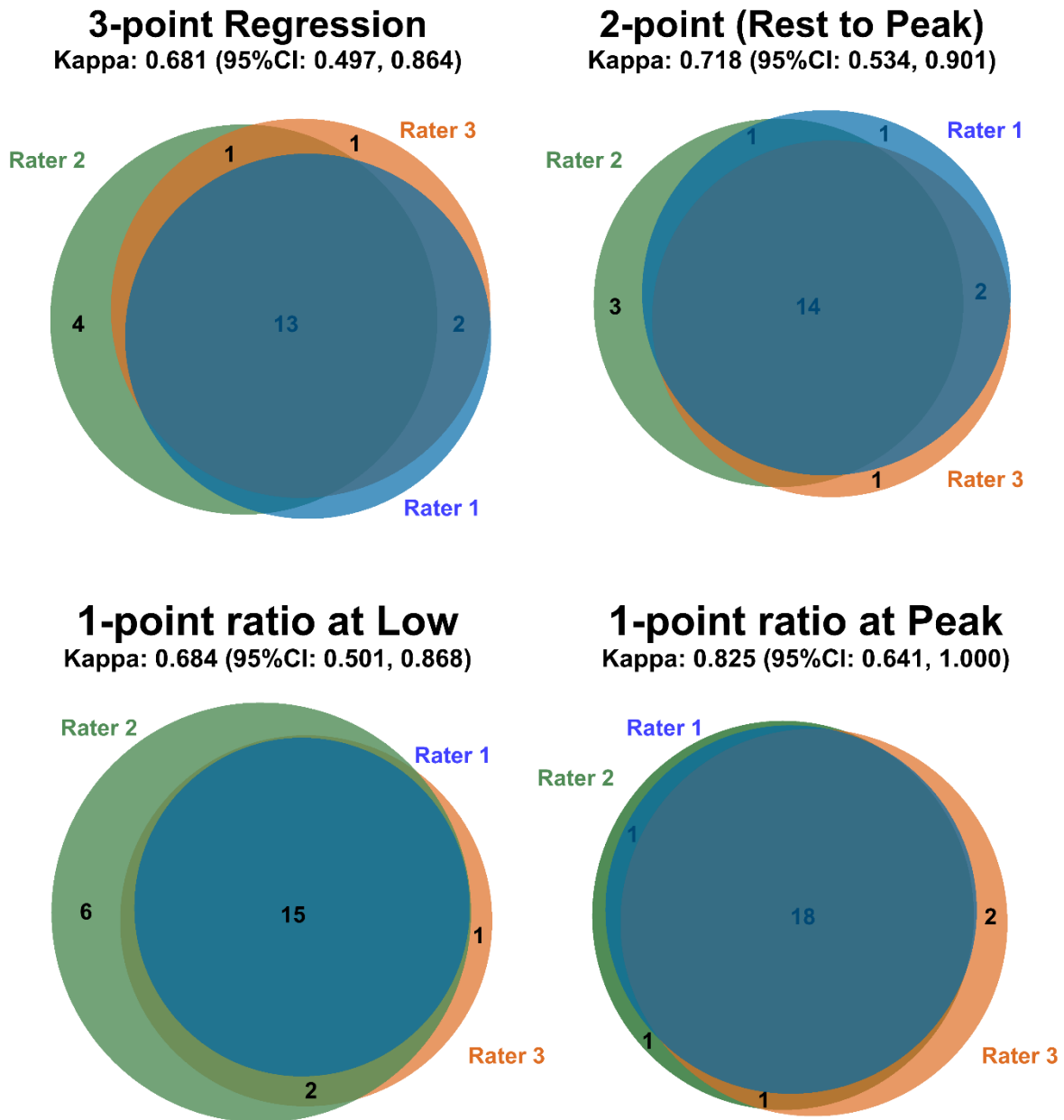
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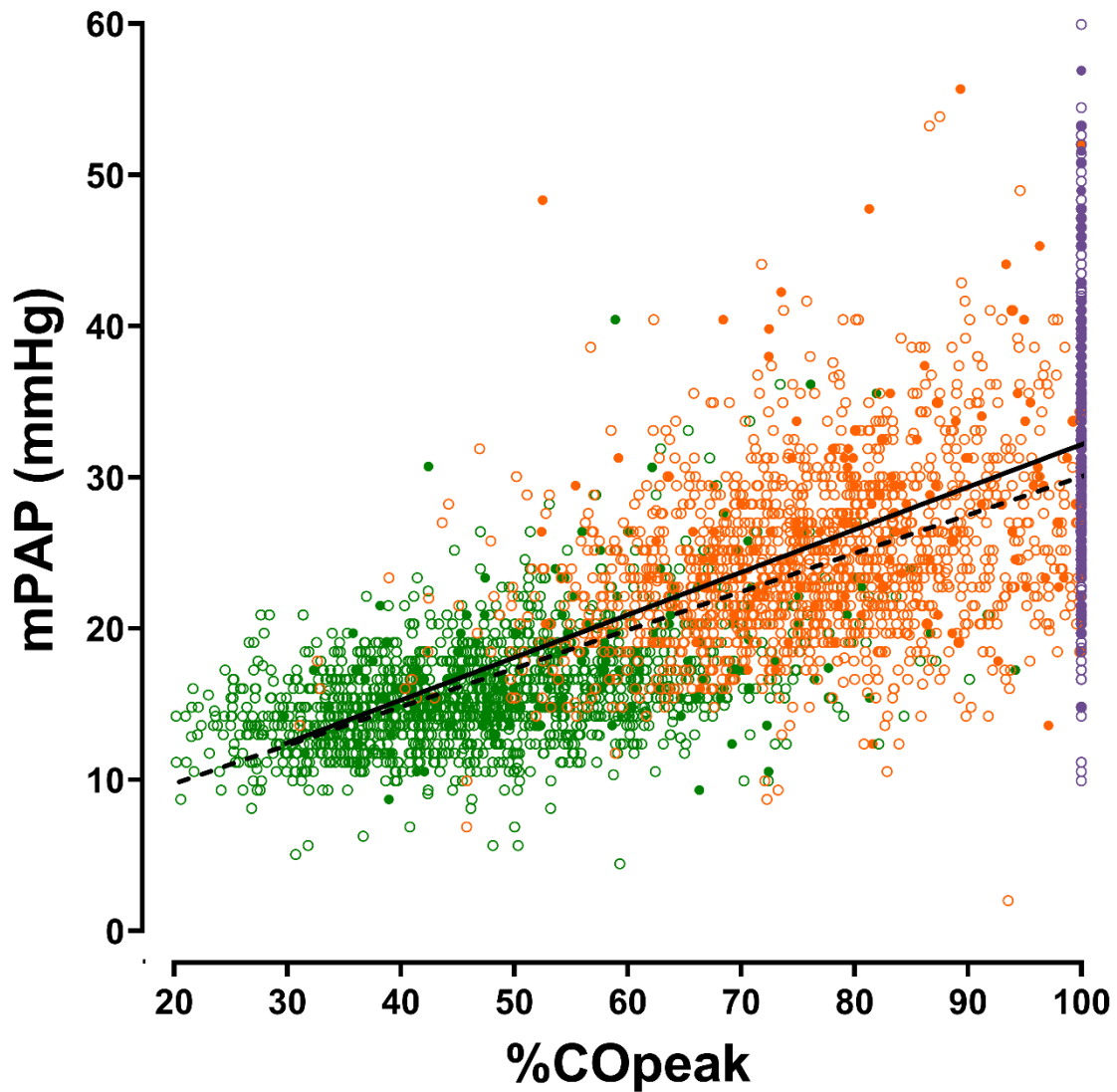
514 **Supplementary Figure 2: Venn diagram for inter-rater agreement in identifying exercise**  
515 **pulmonary hypertension.**



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517 **Supplementary Figure 2:** Inter-rater agreement across the four methods for evaluating the mPAP-CO  
518 relationship. Venn diagrams illustrate diagnostic concordance among three raters. Fleis' kappa values (with 95%  
519 CI) indicate the level of agreement for each method, with the highest reproducibility observed for one-point  
520 mPAP/CO ratio at peak exercise.

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522 **Supplementary Figure 3: Hemodynamic response of mPAP to increasing Cardiac Output during**  
523 **Exercise**



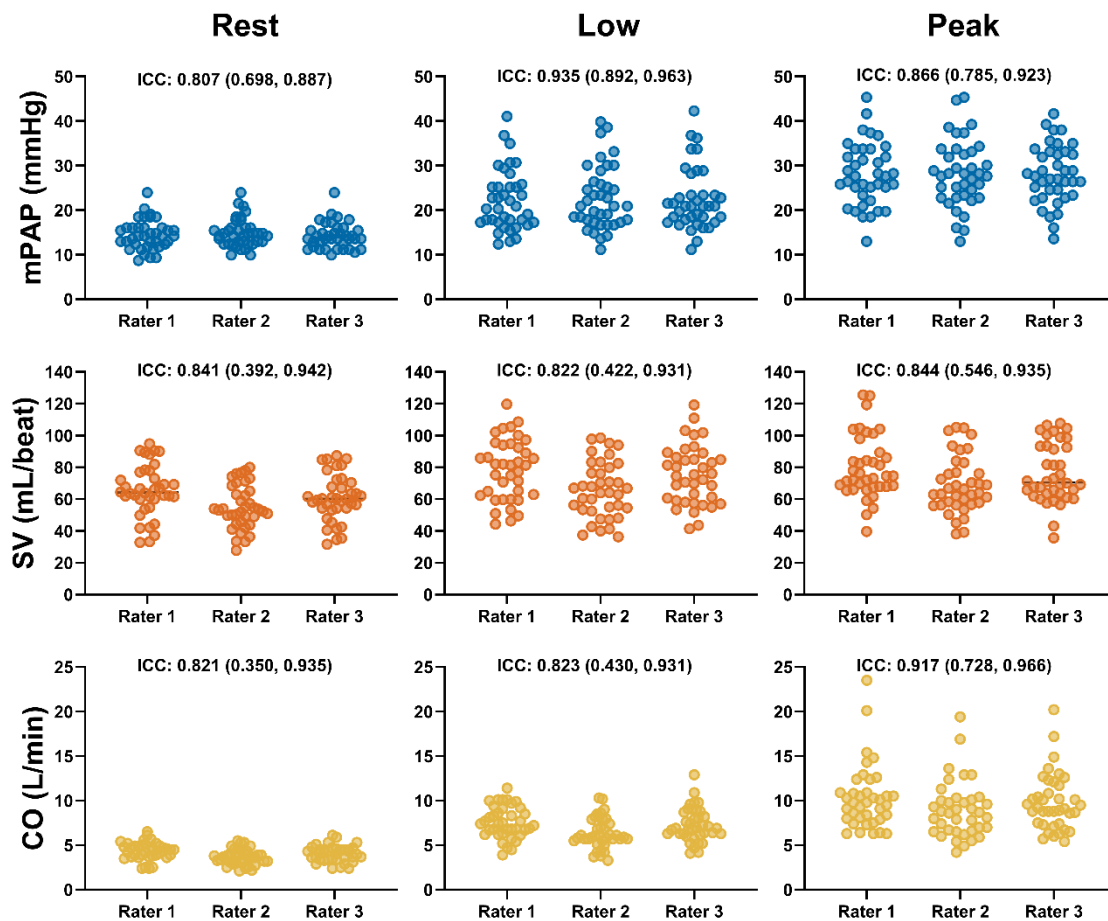
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525 **Supplementary Figure 3:** Individual measurements during exercise are plotted across different stages (green:  
526 rest; orange: low-level exercise; purple: peak exercise). Cardiac output (CO) is expressed as a percentage of peak  
527 cardiac output. A linear relationship between CO and mPAP is observed from rest to low-level exercise (solid line:  
528 individuals with events; dashed line: individuals without events).  
529 CO, cardiac output; mPAP, mean pulmonary artery pressure

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532 **Supplementary Figure 4: Interrater variability expressed as intraclass correlation coefficient (ICC) for**  
533 **the individual components of the mPAP-CO relationship.**

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535

536 *Supplementary Figure 4: Data expressed as intraclass correlation coefficient (ICC) with its corresponding 95%*  
537 *confidence interval.*

538 *Evaluation of the individual components of the mPAP-CO relationship did not reveal any single factor that*  
539 *accounted for the observed differences in ICCs across methods. Slope-based approaches, which rely on multiple*  
540 *measurements each subject to some degree of imprecision, are more susceptible to compounding error. In*  
541 *contrast, one-point mPAP/CO ratios are inherently less affected by cumulative variability*

**Supplementary Table 1.** Cox regression univariate and multivariate models for mPAP/CO methods.

mPAP/CO method	3-point Regression		2-point (Rest to Peak)		1-point ratio at Low		1-point ratio at Peak	
	Hazard ratio	Harrell's C-Index	Hazard ratio	Harrell's C-Index	Hazard ratio	Harrell's C-Index	Hazard ratio	Harrell's C-Index
<b>Univariate models with mPAP/CO (n = 1619, 137 events)</b>								
mPAP/CO method	1.18 (1.13-1.23) p < 0.001	0.696 (0.644-0.748)	1.15 (1.10-1.19) p < 0.001	0.694 (0.643-0.744)	1.52 (1.38-1.67) p < 0.001	0.715 (0.666-0.764)	1.45 (1.32-1.59) p < 0.001	0.720 (0.674-0.766)
<b>Multivariate model A (mPAP/CO, age and cardiorespiratory fitness as a continuous variable; n = 1583, 134 events)</b>								
mPAP/CO method	1.06 (1.00-1.12) p = 0.06	0.752 (0.712-0.793)	1.04 (0.99-1.10) p = 0.14	0.750 (0.709-0.790)	1.21 (1.07-1.38) p = 0.003	0.761 (0.719, 0.802)	1.14 (1.01-1.29) p = 0.04	0.755 (0.715-0.796)
Age, year	1.05 (1.04-1.07) p < 0.001		1.06 (1.04-1.07) p < 0.001		1.05 (1.03-1.07) p < 0.001		1.05 (1.03-1.07) p < 0.001	
VO <sub>2</sub> peak, ml/kg/min	0.96 (0.94-0.99) p = 0.001		0.96 (0.94-0.98) p < 0.001		0.97 (0.95-0.99) p = 0.004		0.96 (0.94-0.99) p = 0.002	
<b>Multivariate model B (mPAP/CO and cardiorespiratory fitness as a categorical variable; n = 1583, 134 events)</b>								
mPAP/CO method	1.15 (1.10-1.20) p < 0.001	0.714 (0.666, 0.763)	1.12 (1.07-1.17) p < 0.001	0.712 (0.664-0.761)	1.44 (1.30-1.60) p < 0.001	0.729 (0.682-0.777)	1.37 (1.24-1.51) p < 0.001	0.724 (0.678-0.770)

VO <sub>2</sub> peak (< 80% versus ≥ 80 of predicted)	2.33 (1.54-3.45) p < 0.001		2.33 (1.56-3.57) p < 0.001		2.13 (1.41-3.23) p < 0.001		2.13 (1.43-3.23) p < 0.001	
<b>Multivariate model C (mPAP/CO, cardiorespiratory fitness as a categorical variable and intensity of effort; n = 1350, 97 events)</b>								
mPAP/CO method	1.18 (1.11-1.27) p < 0.001		1.14 (1.07-1.21) p < 0.001		1.42 (1.25-1.60) p < 0.001		1.34 (1.19-1.51) p < 0.001	
VO <sub>2</sub> peak (< 80% versus ≥ 80 of predicted)	2.15 (1.40-3.31) p < 0.001	0.716 (0.661-0.771)	2.18 (1.41-3.36) p < 0.001	0.714 (0.660-0.768)	2.06 (1.34-3.18) p = 0.001	0.727 (0.670-0.785)	2.07 (1.34-3.20) p = 0.001	0.723 (0.668-0.778)
RERpeak (< 1.05 versus ≥ 1.05)	1.75 (1.14-2.70) p = 0.011		1.78 (1.15-2.74) p = 0.001		1.64 (1.06-2.54) p = 0.026		1.63 (1.05-2.53) p = 0.029	

**Supplementary Table 1:** Data expressed as Hazard Ratio or Harrell's C-Index, and 95% confidence interval. Univariable and multivariable Cox models are shown for the four mPAP/CO methods. Multivariable model A adjusts for age and cardiorespiratory fitness (VO<sub>2</sub>peak, continuous); model B adjusts for VO<sub>2</sub>peak as a categorical variable (<80% vs. ≥ 80% predicted); and model C additionally includes effort level (RERpeak <1.05 vs ≥ 1.05).

In the univariable Cox regression, all four mPAP/CO methods were individually associated with clinical outcomes. In multivariable model A, only the one-point mPAP/CO ratio at low and peak remained significant, while the 3-point regression and 2-point (rest to peak) did not. Age and VO<sub>2</sub>peak remained independent predictors. In multivariable model B (VO<sub>2</sub>peak categorical), all models remained statistically significant. In multivariable model C (VO<sub>2</sub>peak and intensity of effort), all models remained statistically significant.

**Supplementary Table 2:** Multivariable logistic regression models evaluating the interaction between resting mPAP and the three-point regression method (model A), two-point slope (model B), the one-point mPAP/CO ratio at low exercise (model C), the one-point mPAP/CO ratio at peak exercise (model D).

Variable	$\beta$	Sig. (p)	Exp(B) (95% CI)	Model
3-point Regression	0.586	<0.001	1.796 (1.397–2.310)	Model A
mPAP (rest)	0.223	<0.001	1.250 (1.158–1.349)	Model A
Interaction (mPAP rest $\times$ 3-point Regression)	-0.023	<0.001	0.978 (0.965–0.991)	Model A
<hr/>				
2-point (Rest to Peak)	0.469	<0.001	1.599 (1.286–1.989)	Model B
mPAP (rest)	0.208	<0.001	1.231 (1.146–1.321)	Model B
Interaction (mPAP rest $\times$ 2-point (Rest to Peak))	-0.018	0.002	0.982 (0.971–0.994)	Model B
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1-point mPAP/CO ratio at low	0.764	<0.001	2.147 (1.393–3.308)	Model C
mPAP (rest)	0.207	<0.001	1.230 (1.095–1.381)	Model C
Interaction (mPAP (rest) $\times$ 1-point mPAP/CO ratio at low)	-0.025	0.027	0.975 (0.954–0.997)	Model C
<hr/>				
1-point mPAP/CO ratio at peak	0.911	<0.001	2.488 (1.593–3.884)	Model D
mPAP (rest)	0.23	<0.001	1.258 (1.116–1.418)	Model D
Interaction (mPAP (rest) $\times$ 1-point mPAP/CO ratio at peak)	-0.031	0.008	0.969 (0.947–0.992)	Model D

**Supplementary Table 2:** Results are presented as beta coefficients ( $\beta$ ), which estimate the change in the Log odds of the outcome for a one-unit increase in the predictor. Odds ratios (Exp(B)) with corresponding 95% confidence intervals (95% CI) are also reported.

The one-point mPAP/CO ratio at peak exercise showed a stronger association with outcome compared to the three-point slope and the two-point slope ( $\beta = 0.911$  vs.  $\beta = 0.469$  vs.  $\beta = 0.586$ , respectively). When mPAP at rest was higher, the additional risk contributed by each method was slightly lower, as reflected by the negative interaction terms (model A:  $\beta = -0.023$ , model B:  $-0.018$ , model C:  $-0.025$ , model D:  $-0.031$ ). However, the magnitude of the interaction effect was low and comparable between methods, suggesting that resting mPAP only modestly attenuates the prognostic value of the different methods.

CO, cardiac output; mPAP, mean pulmonary artery pressure

**Supplementary Table 3: Correlation and collinearity diagnostic for mPAP/CO methods**

Variable	Correlation matrix (mPAP at rest)	Variance Inflation Factors (VIF)
3-point Regression	0.34	1.13
2-point (Rest to Peak)	0.32	1.12
1-point ratio at Low	0.55	1.44
1-point ratio at Peak	0.54	1.42

Supplementary Table 3: This table shows the pairwise correlation between mPAP and the four mPAP/CO methods, as well as their corresponding Variance Inflation Factors (VIF). Correlations were modest ( $r = 0.32$ – $0.55$ ), and all VIF values were low ( $1.11$ – $1.44$ ), indicating no evidence of problematic multicollinearity.

*CO, cardiac output; mPAP, mean pulmonary artery pressure*